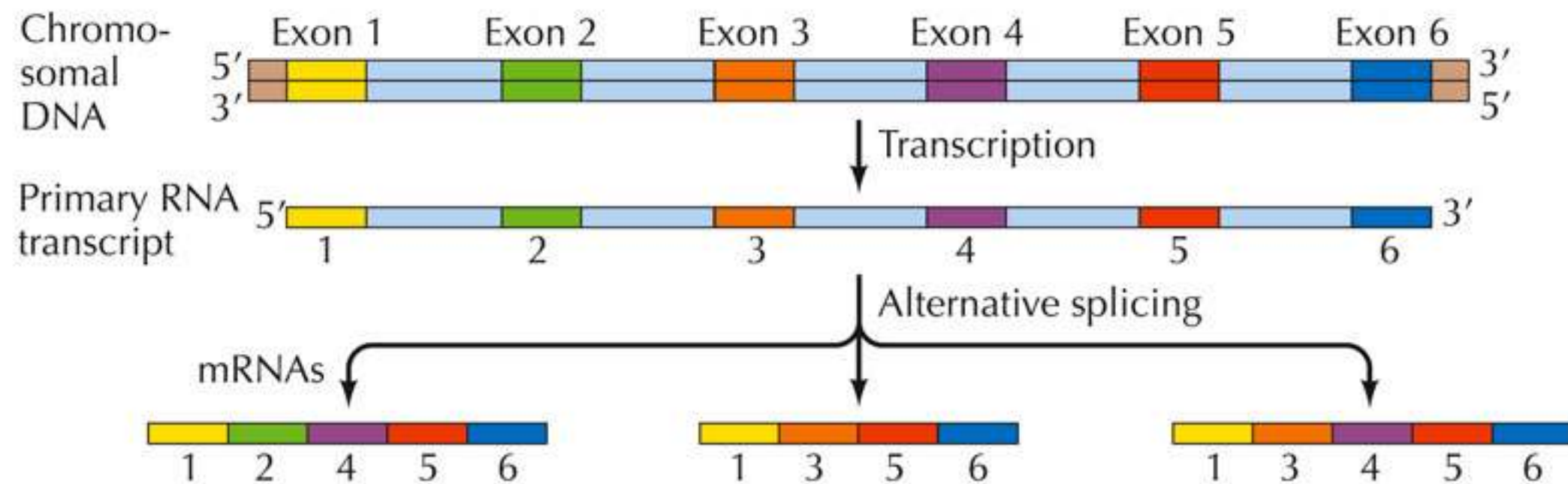


Differential expression

Alejandro Reyes
T: @areyesq89

Workshop on Transcriptomics
September 13th, 2017

Overview of exons, genes and transcripts



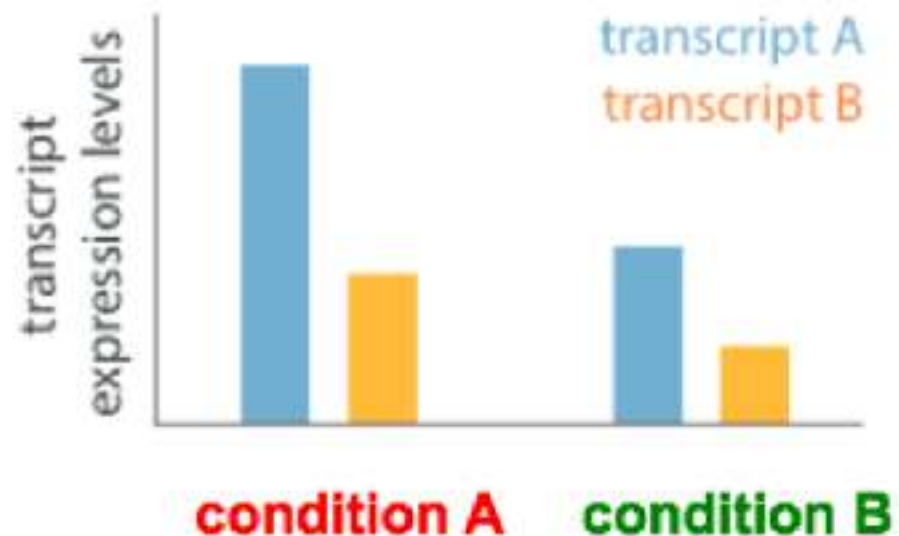
What is your biological question?

Given a gene, test for:

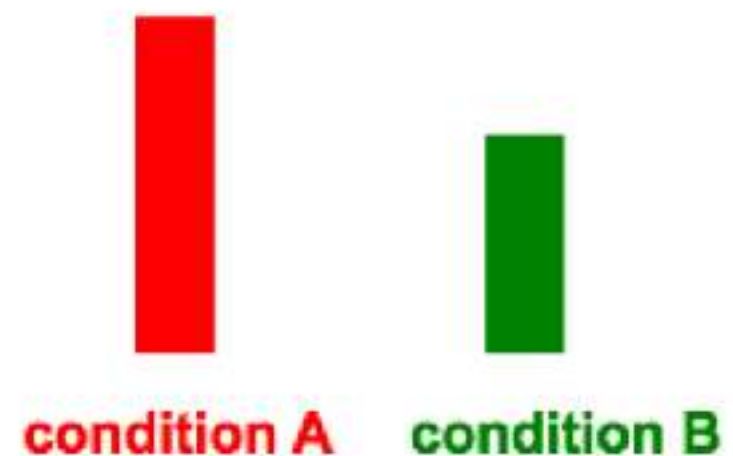
- Whether transcripts levels change between conditions? (differential exon usage, DGE)
- Whether transcript isoform proportions change between conditions? (differential transcript usage, DTU)
- Whether individuals exons are differentially used? (differential exon usage, DEU)

Differential problems: DGE

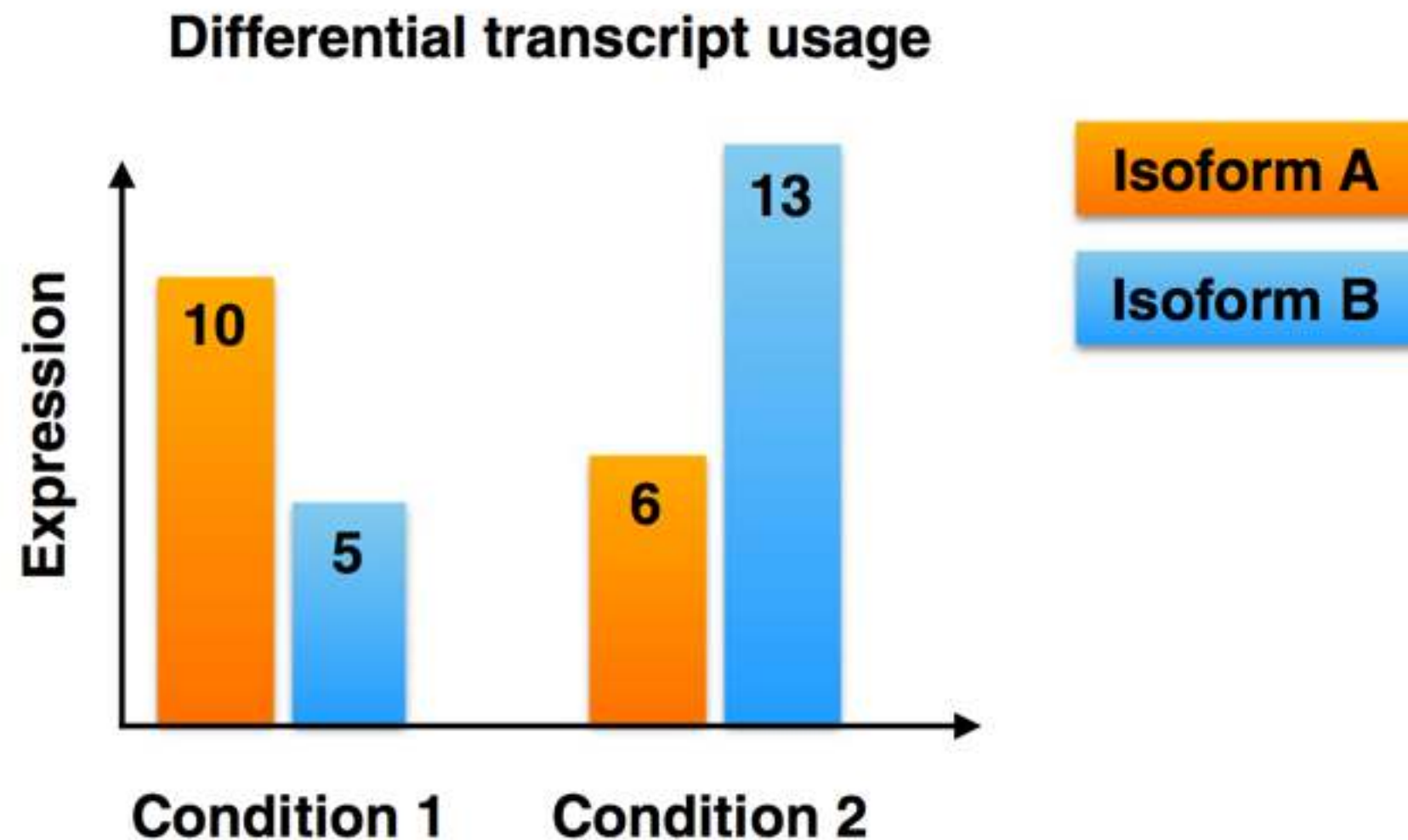
Differential transcript
expression (DTE)



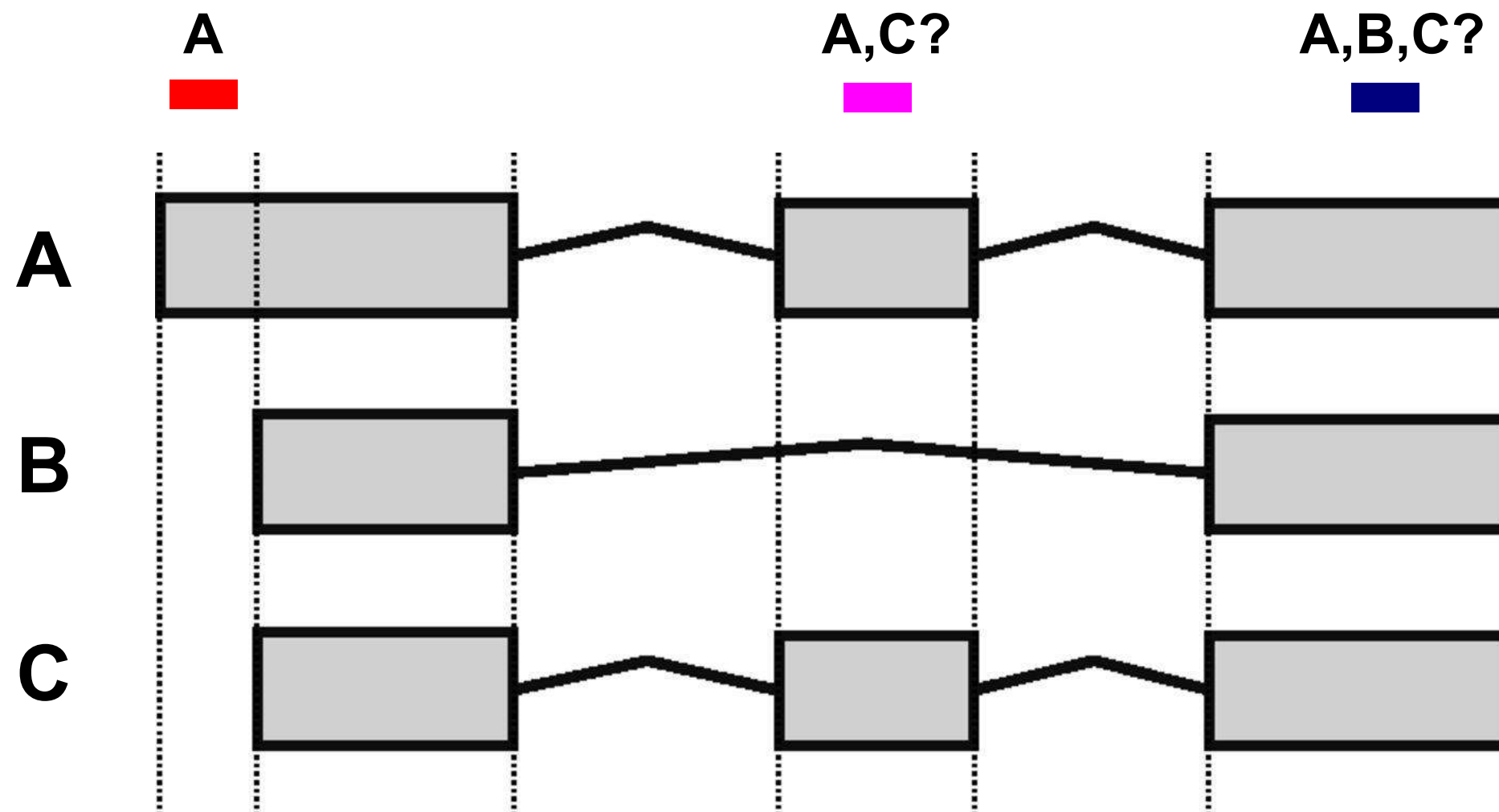
Differential gene
expression (DGE)



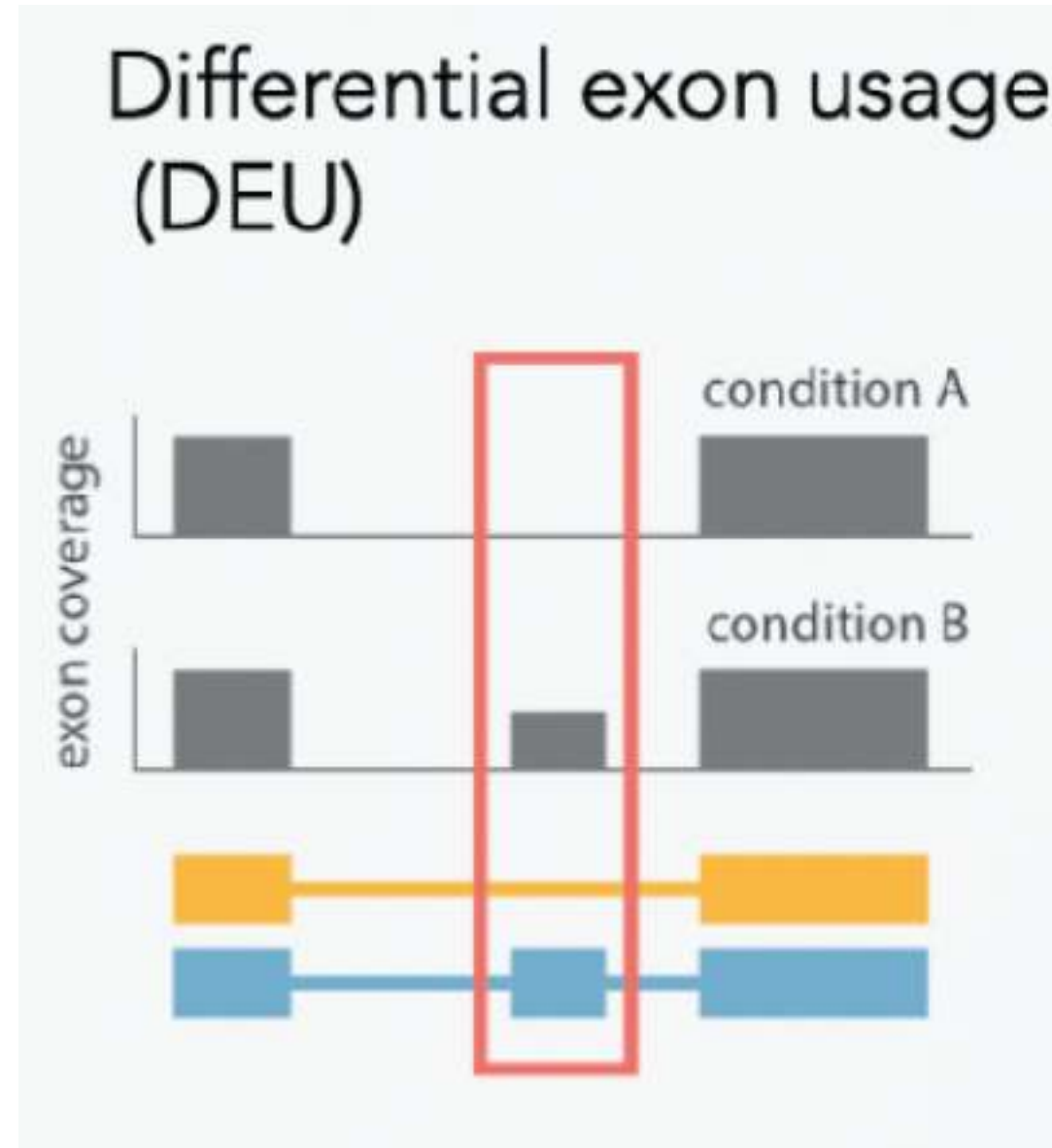
Differential problems: DTU



Unambiguous assignment of reads to single transcripts



Differential problems: DEU

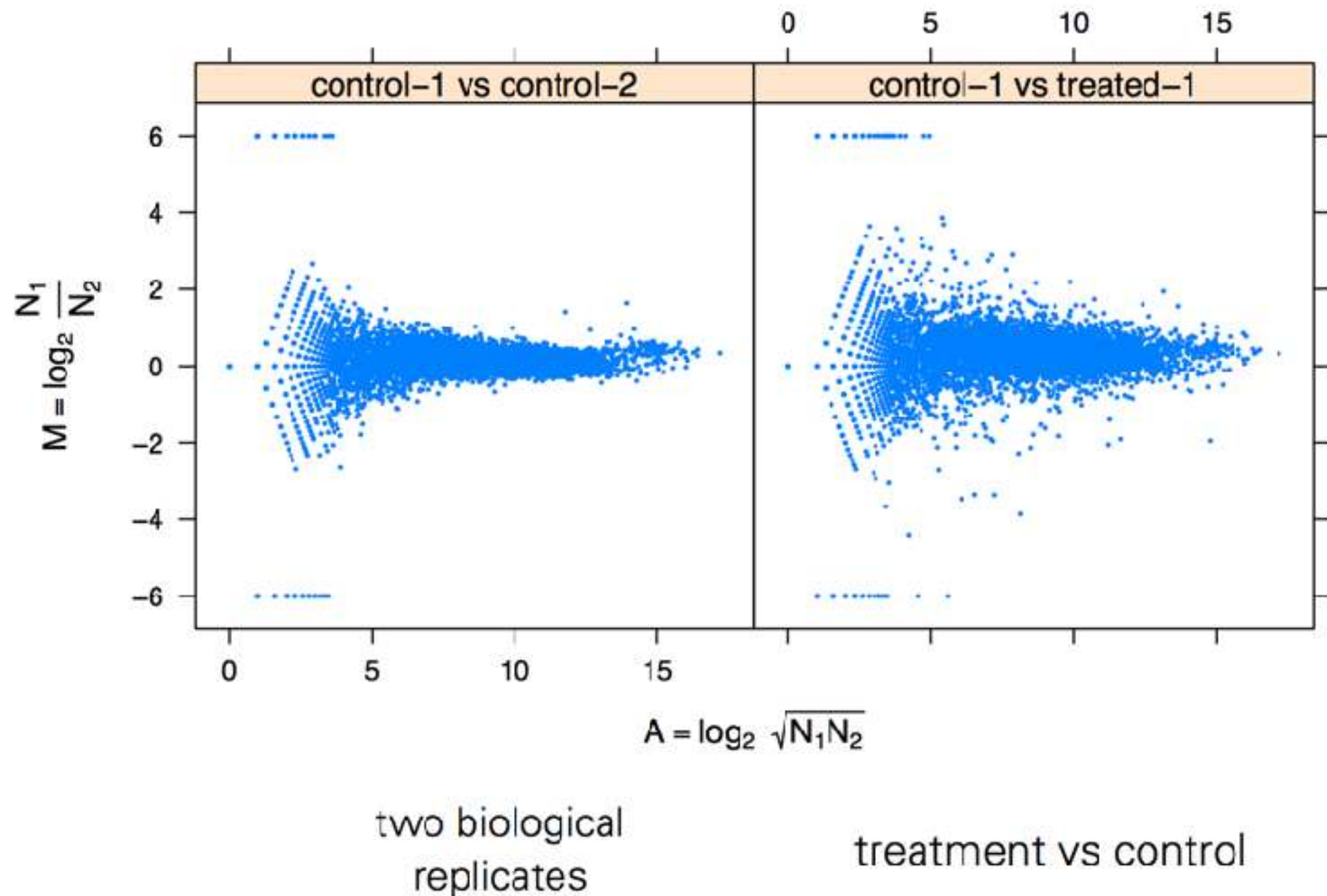


What is your biological question?

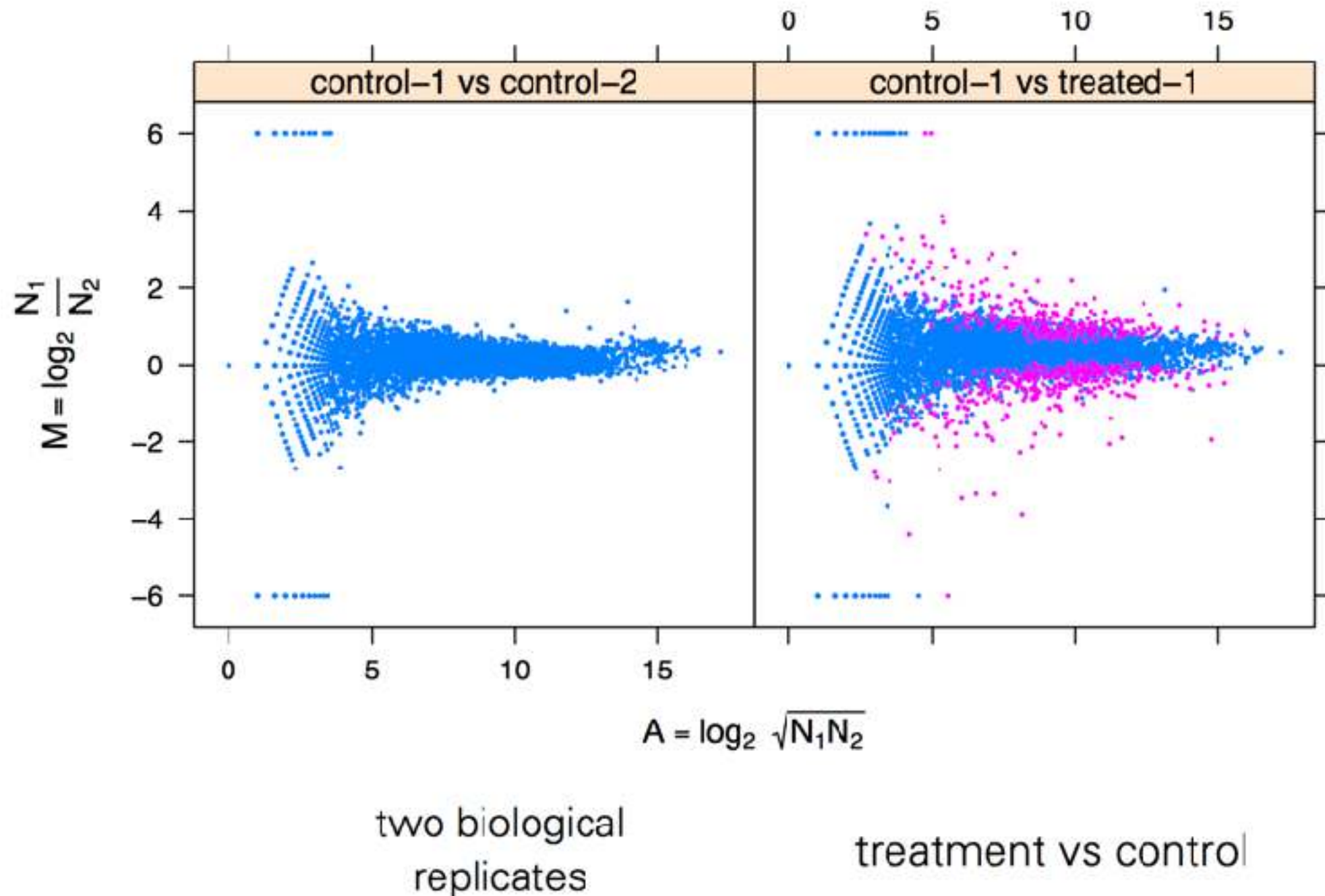
Given a gene, test for:

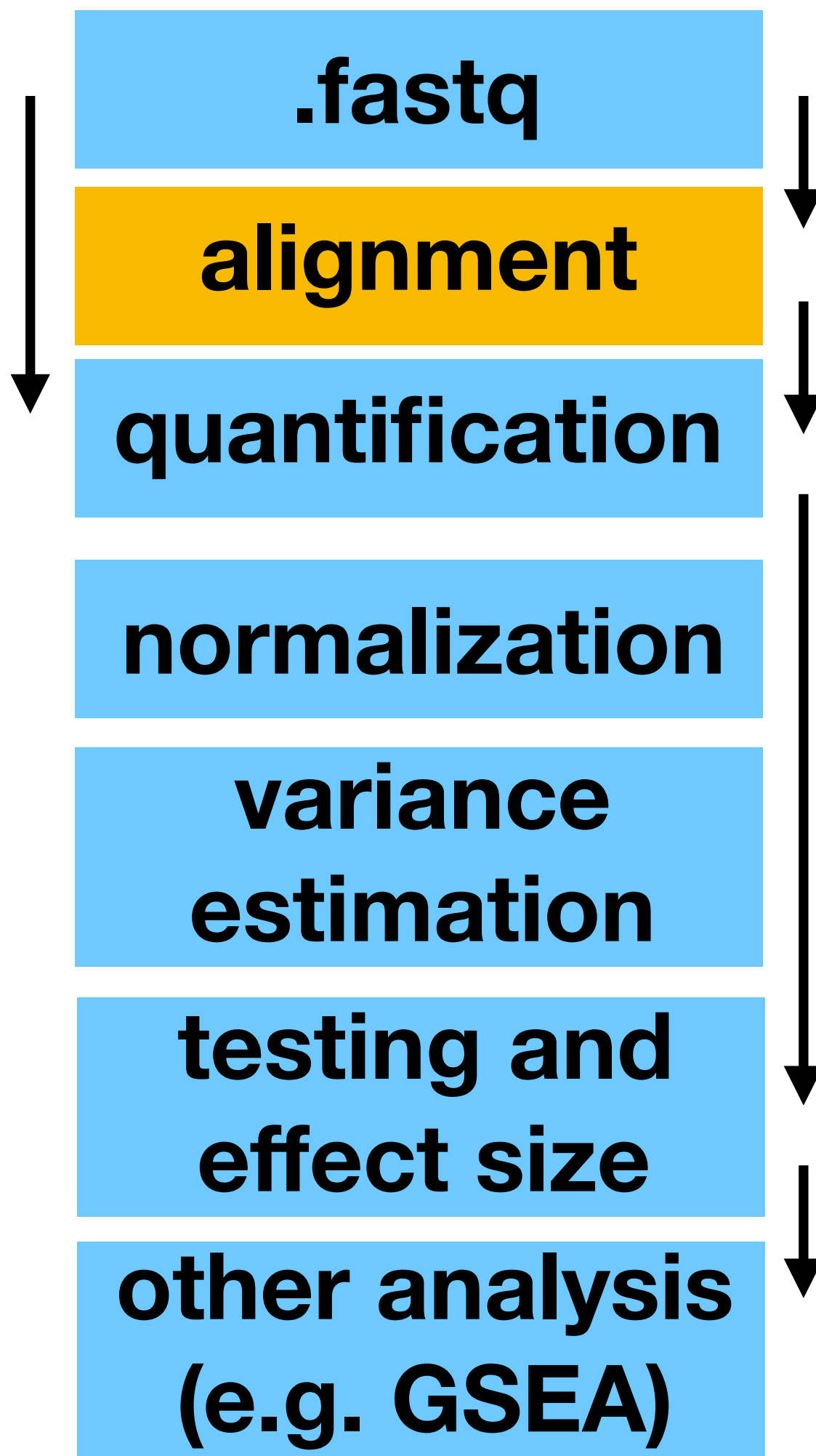
- **Whether transcripts levels change between conditions? (differential gene expression, DGE)**
- Whether transcript isoform proportions change between conditions? (differential transcript usage, DTU)
- Whether individuals exons are differentially used? (differential exon usage, DEU)

Differential problems: DGE



Differential problems: DGE





salmon,
kallisto,
...

.fastq

alignment

quantification

normalization

**variance
estimation**

**testing and
effect size**

**other analysis
(e.g. GSEA)**



STAR, GSNAP, ...



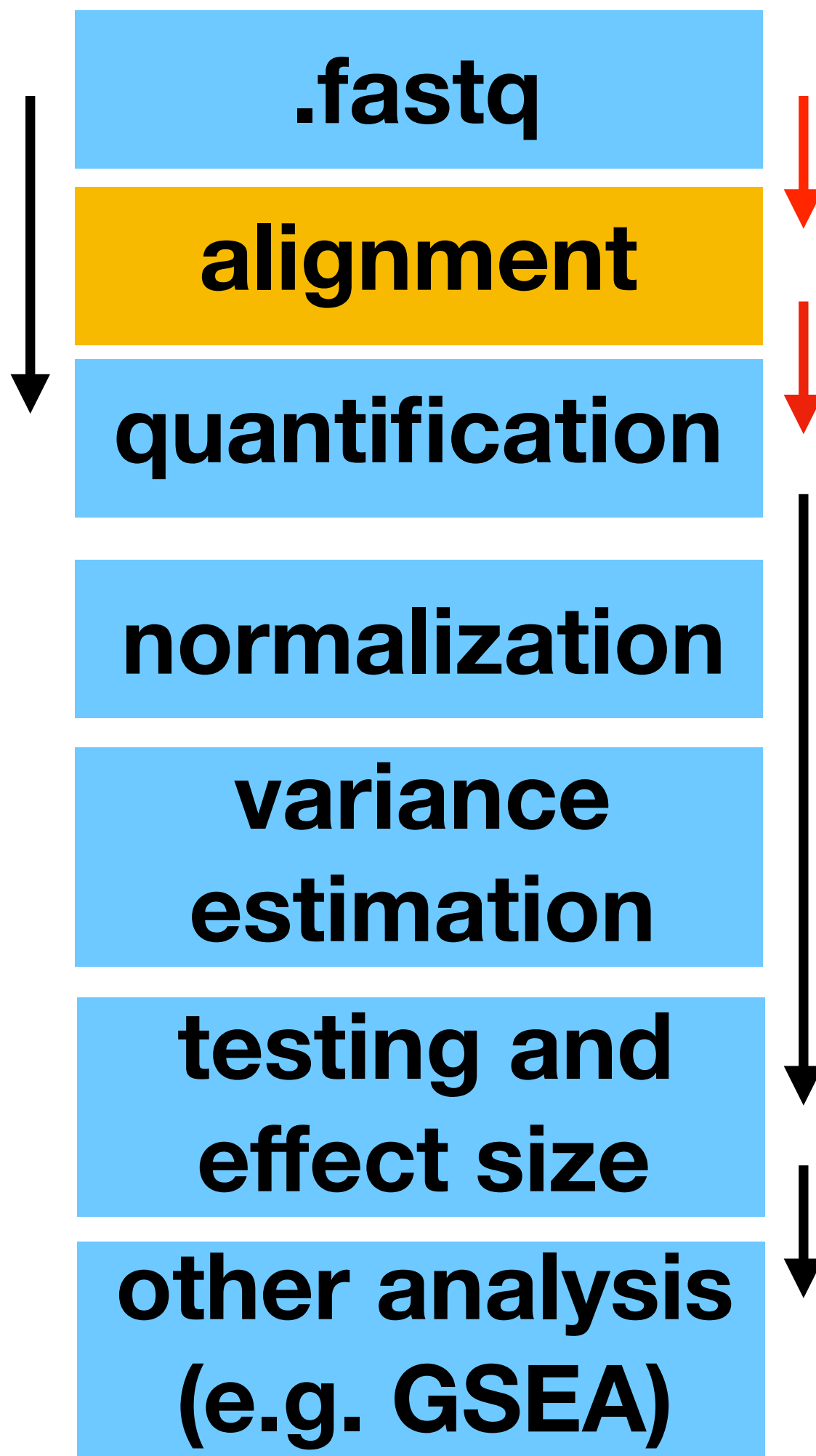
ht-seq, featureCounts,



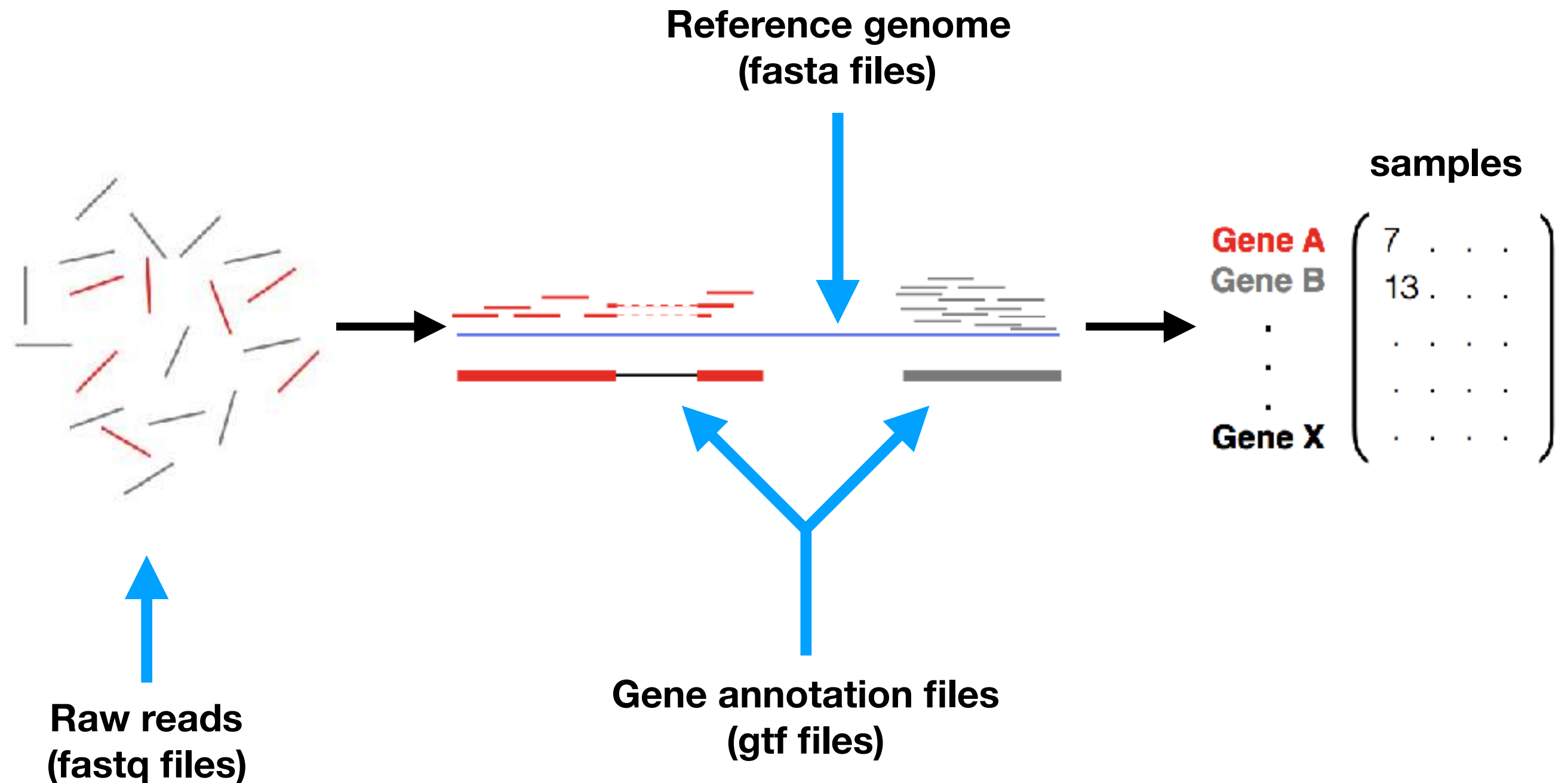
DESeq2,
edgeR,
limma-voom
NOISeq,
sleuth,...



goseq, roast, ...

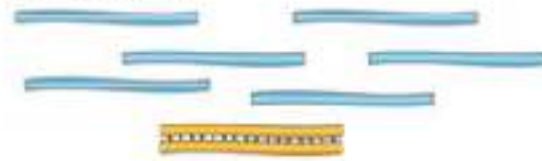


Alignment-based abundance estimation workflow



a Data generation

① mRNA or total RNA



② Remove contaminant DNA



③ Fragment RNA



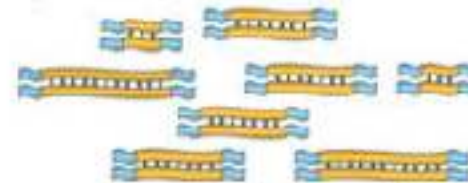
Remove rRNA?
Select mRNA?

④ Reverse transcribe
into cDNA



Strand-specific RNA-seq?

⑤ Ligate sequence adaptors



PCR amplification?

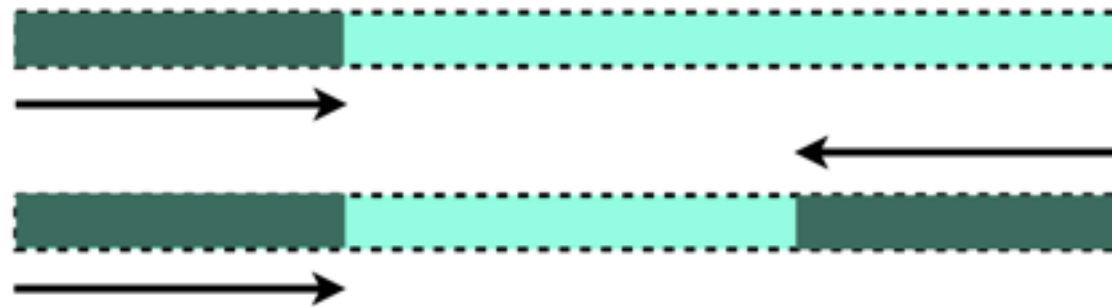
⑥ Select a range of sizes



⑦ Sequence cDNA ends

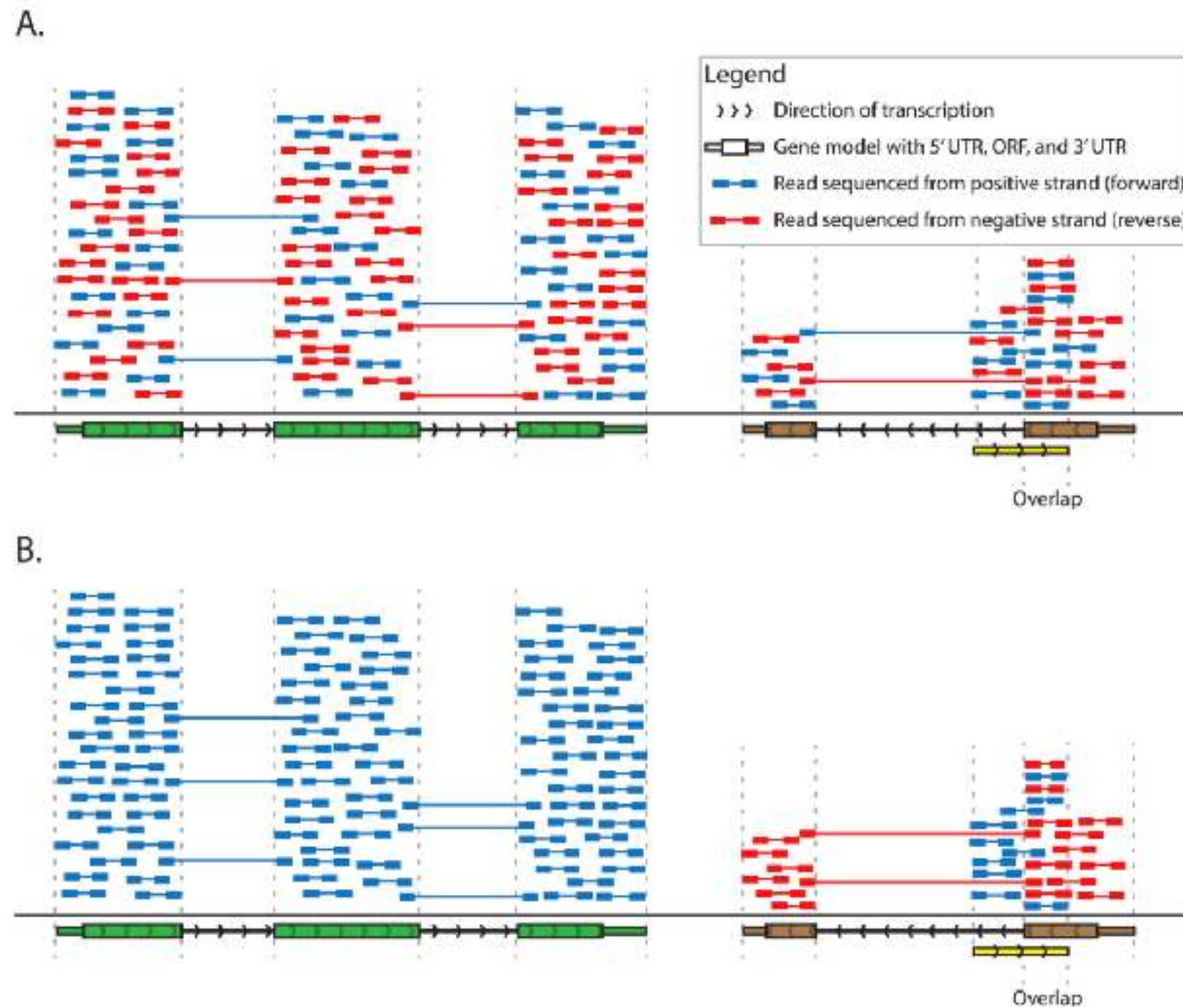


Are my data single-end or paired-end?



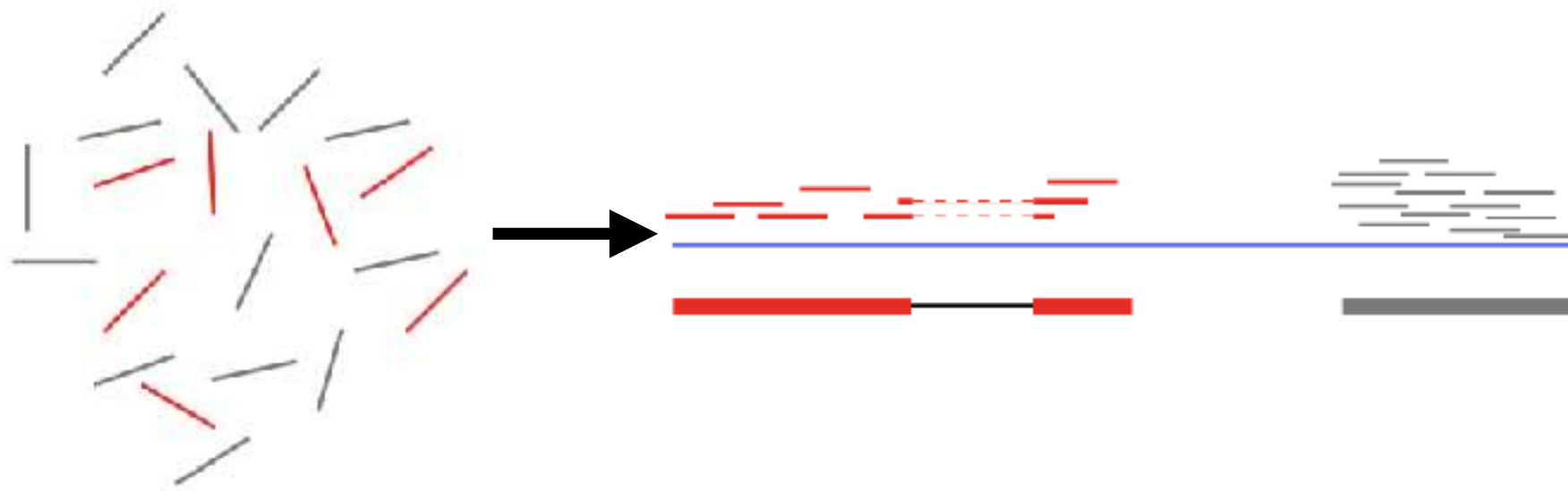
Tip: Look at the number of files per sample

Are my data strand specific?



Tip: Visualize mapped reads in a genome browser (e.g. IGV)

Alignment



**Need a splice aware aligner
(e.g. STAR, GSNAP, ...)**

NATURE METHODS | ANALYSIS



Simulation-based comprehensive benchmarking of RNA-seq aligners

Giacomo Baruzzo, Katharina E Hayer, Eun Ji Kim, Barbara Di Camillo, Garret A FitzGerald & Gregory R Grant

NATURE METHODS | ANALYSIS [OPEN](#)



Systematic evaluation of spliced alignment programs for RNA-seq data

Pär G Engström, Tamara Steijger, Botond Sipos, Gregory R Grant, André Kahles, The RGASP Consortium, Gunnar Rätsch, Nick Goldman, Tim J Hubbard, Jennifer Harrow, Roderic Gulgó & Paul Bertone

Example using STAR: index generation

```
$ STAR --runThreadN 24 \  
--runMode genomeGenerate \  
--genomeDir my_genome \  
--genomeFastaFiles my_genome.fa \  
--sjdbGTFfile my_genes.gtf \  
--sjdbOverhang 99
```

number of threads

output folder -
name according
to genome

reference
genome

gene
annotation
file

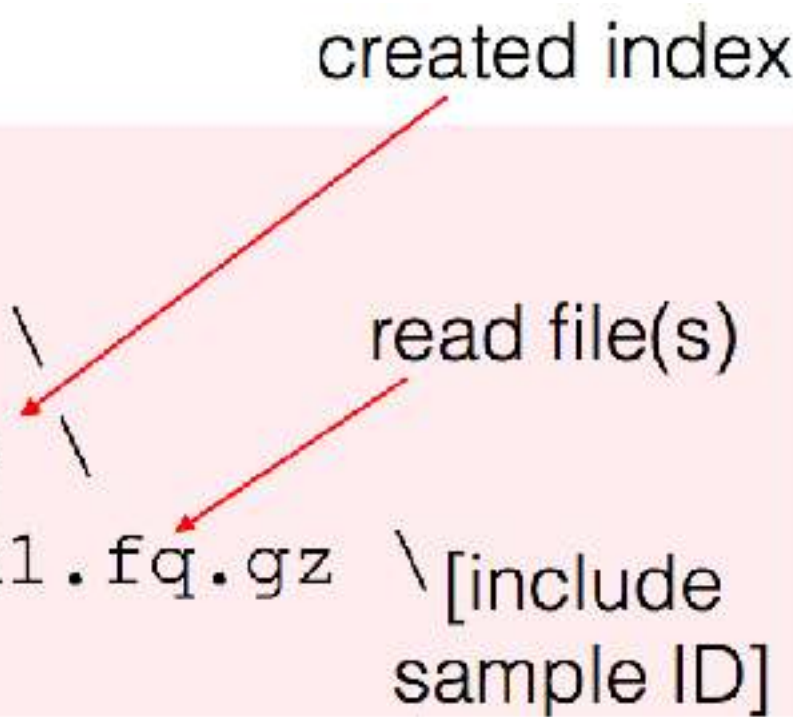
read length - 1

Example using STAR: alignment

```
$ STAR --runThreadN 24 \  
      --runMode alignReads \  
      --genomeDir my_genome \  
      --readFilesIn S1_read1.fq.gz \  
                   S1_read2.fq.gz \  
                   \[include  
                   sample ID]
```

created index

read file(s)

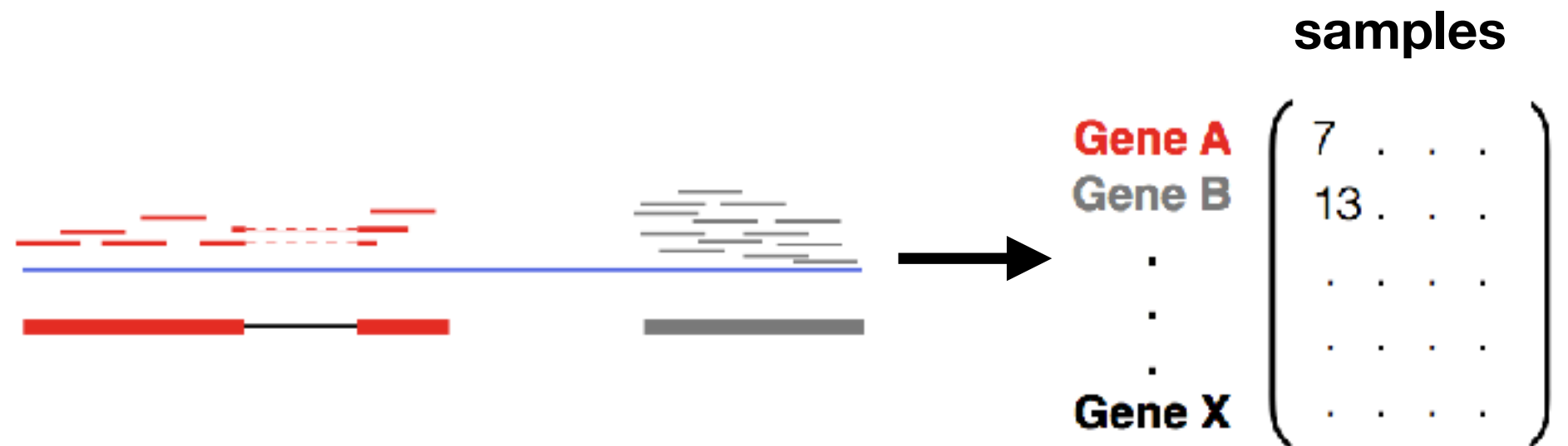


Aligner output: BAM files

[illegible]

read name **flag** **chr** **pos** **mapq** **CIGAR**

Alignment-based abundance estimation workflow



format bam \
--order=pos \
--stranded=no \
--type=exon \
--idattr=gene_id \
--mode=union \
aligned.bam \
my_genes.gtf

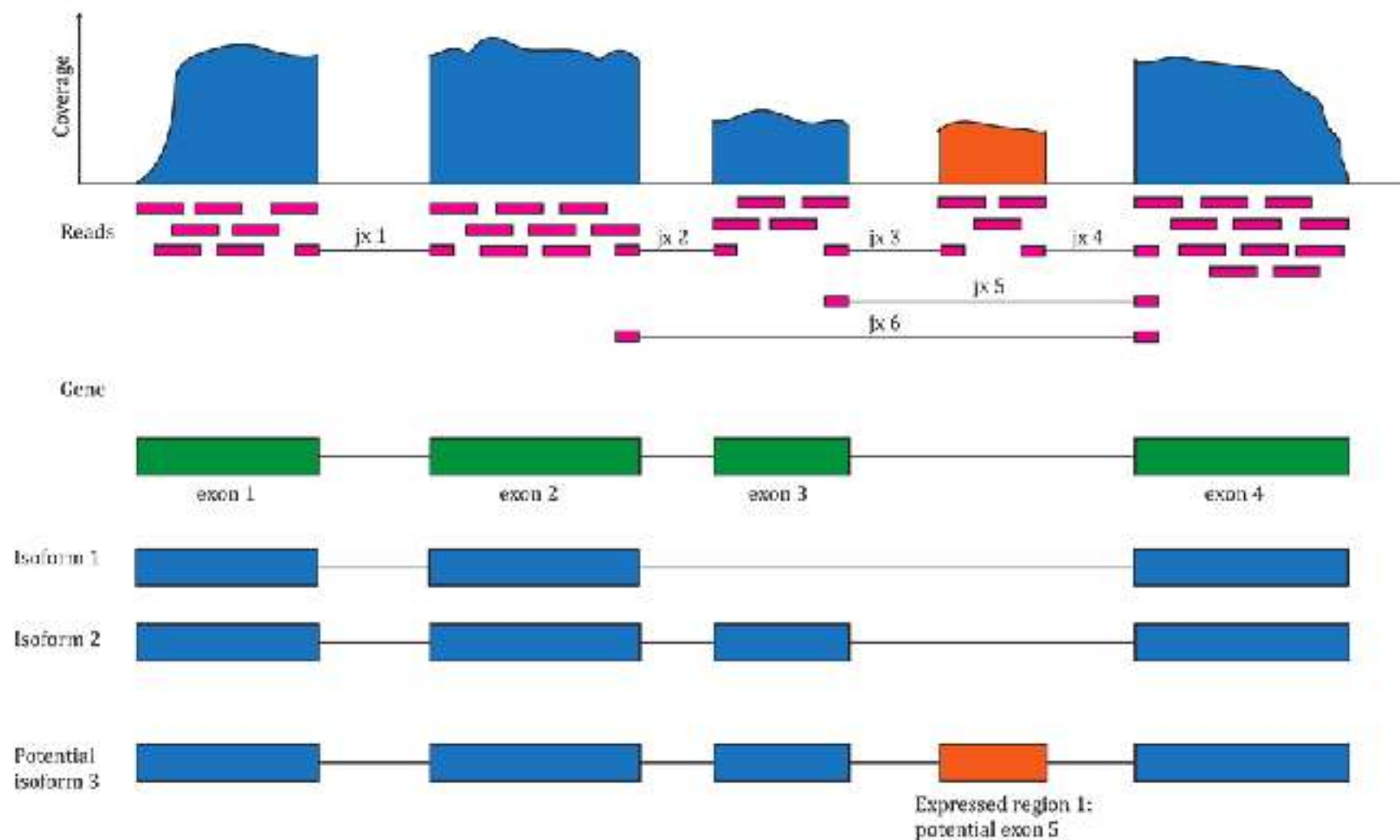
default=yes
check you
GTF file!

protein_coding exon 5139815 5141712 . - . gene_id "FBgn0020621"; transcript_id
Btr0112897"; exon_number "10"; gene_name "Pkn"; gene_biotype "protein_coding"; transcript_name "Pkn-RG"
on_id "FBgn0020621:1";

mapped reads
gene
annotation file (GTF)

recount2

70,000 human RNA-seq samples
already processed





.fastq

alignment

quantification

normalization

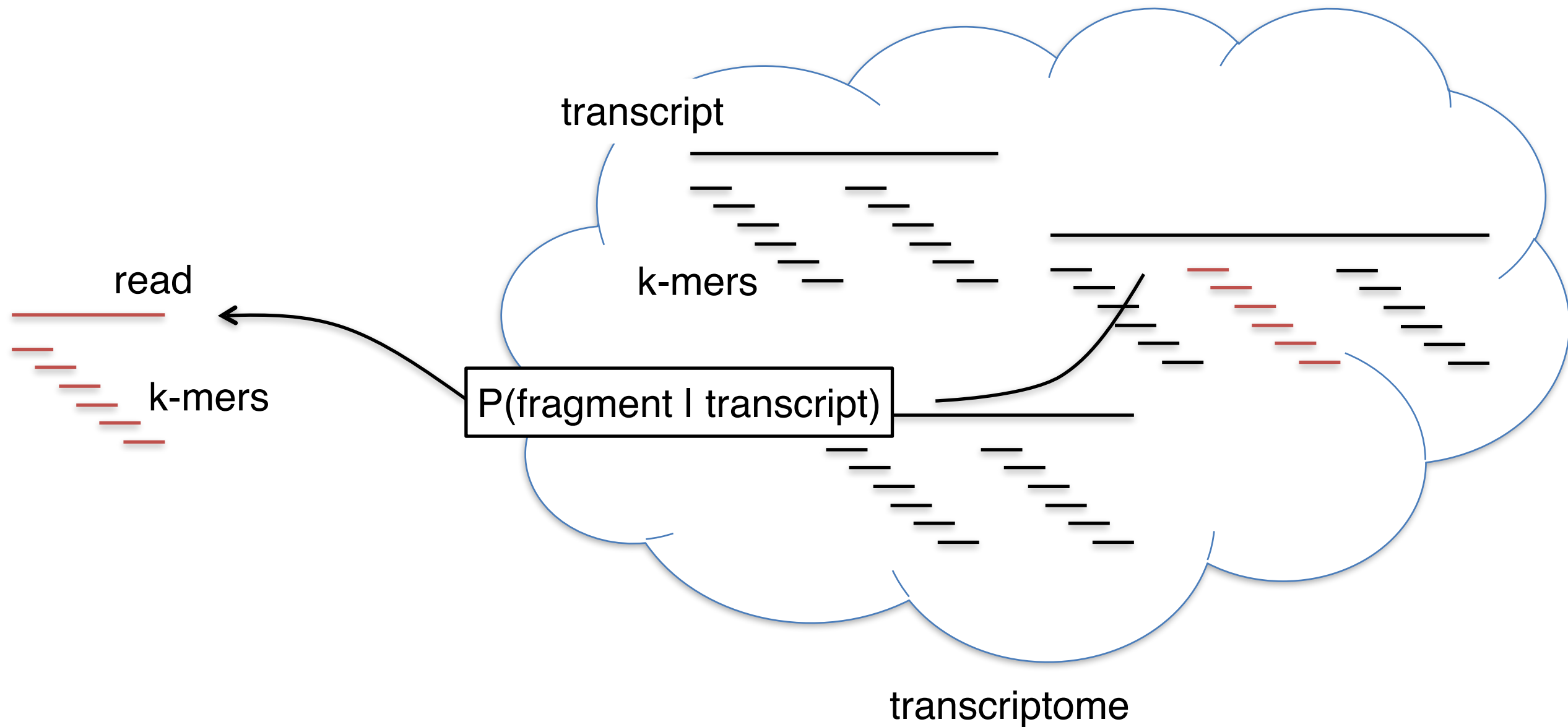
**variance
estimation**

**testing and
effect size**

**other analysis
(e.g. GSEA)**

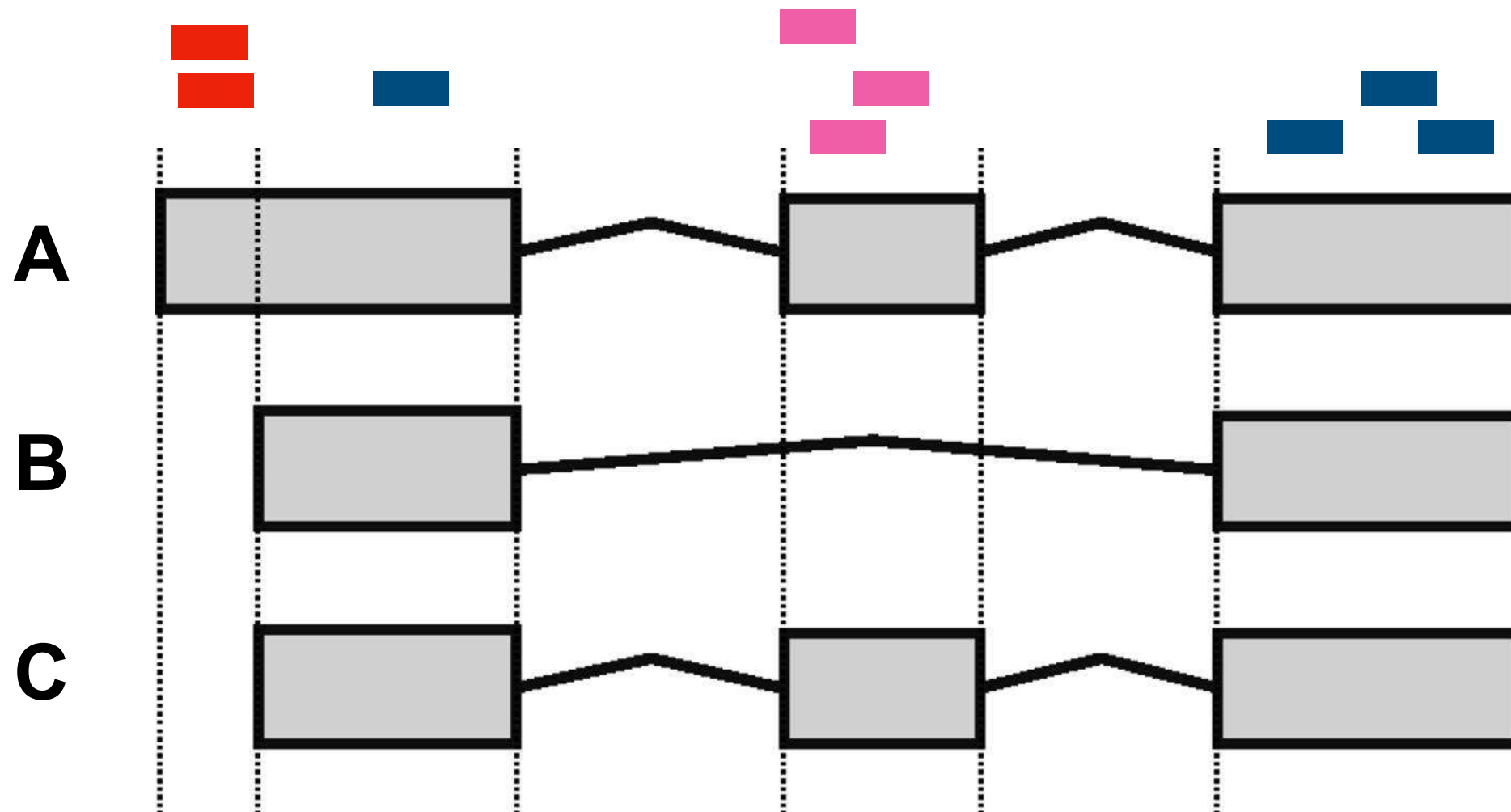


Alignment-free abundance estimation workflows



Extremely fast compared to genome alignments!

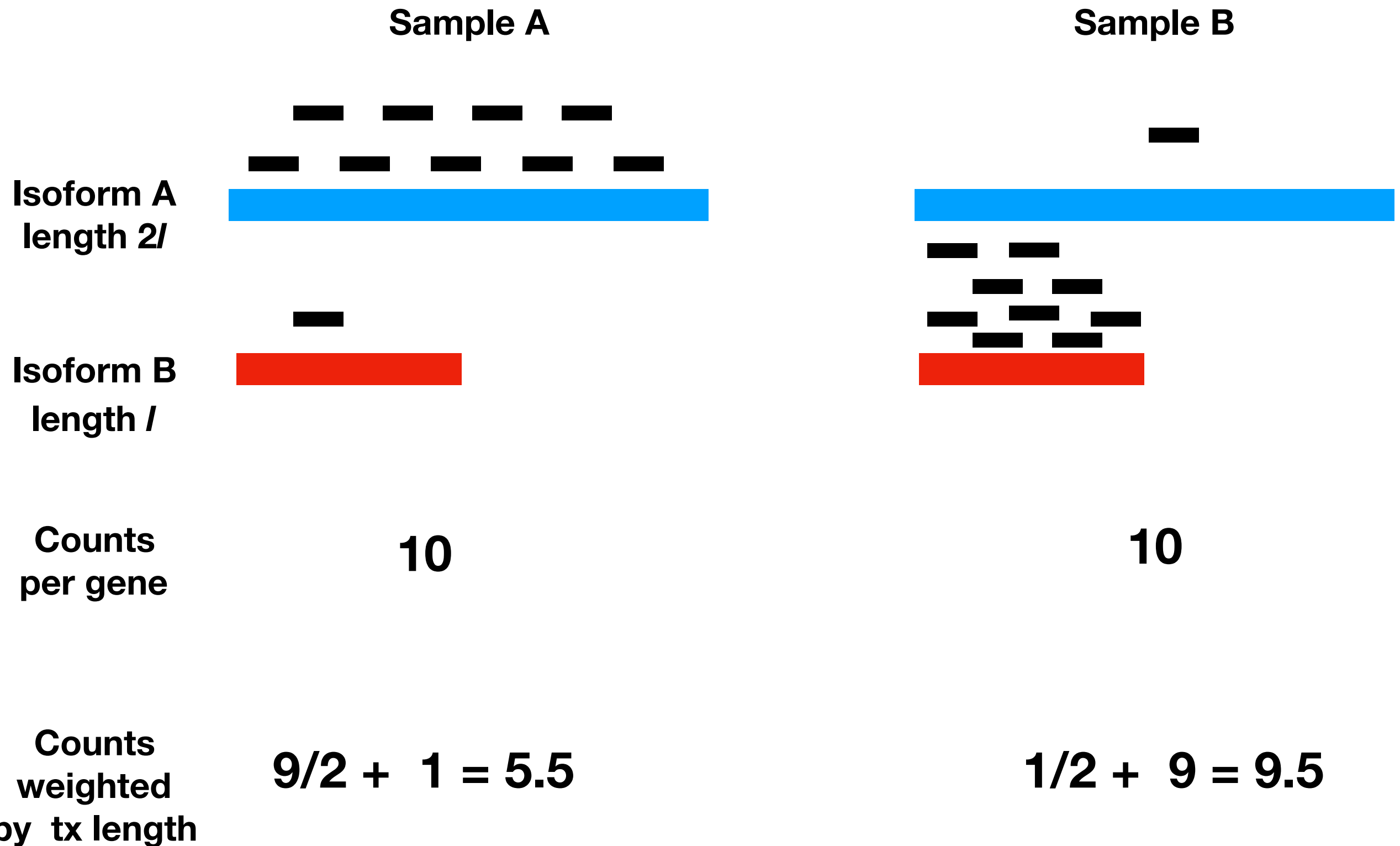
Alignment-free abundance estimation workflows



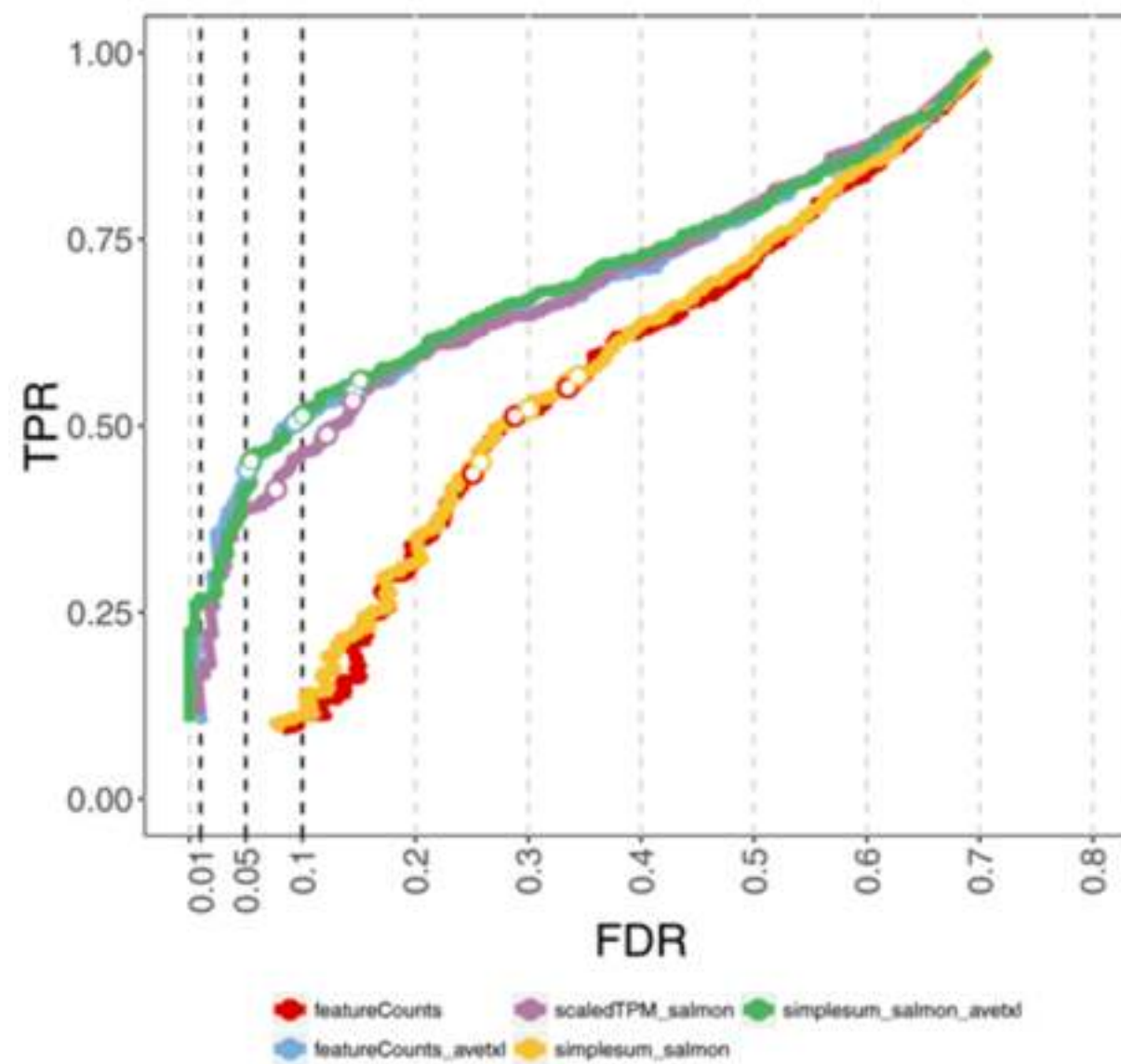
Equivalence classes: $\{A\} = 2$, $\{A, C\} = 3$, $\{A, B, C\} = 4$

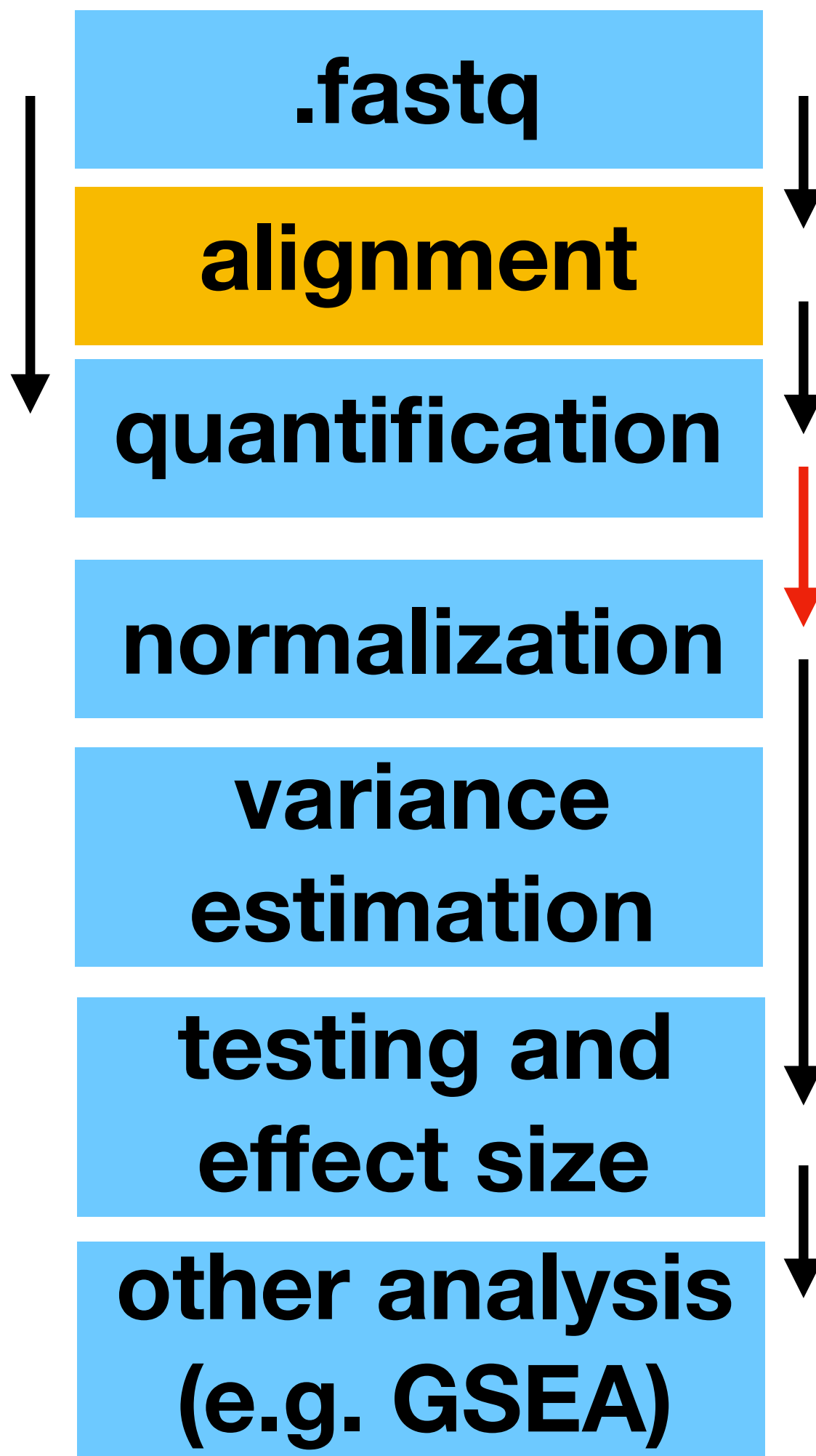
Objective: estimate the transcript abundances that best explain the observed counts for the reference classes.

Why is it important to consider transcript length information? (hypothetical example)



Considering average transcript lengths improves estimates of DGE





$$\begin{array}{l}
 \text{Gene A} \\
 \text{Gene B} \\
 \vdots \\
 \text{Gene X}
 \end{array}
 \begin{pmatrix}
 7 & . & . & . \\
 13 & . & . & . \\
 . & . & . & . \\
 . & . & . & . \\
 . & . & . & .
 \end{pmatrix}$$

Sequencing depth

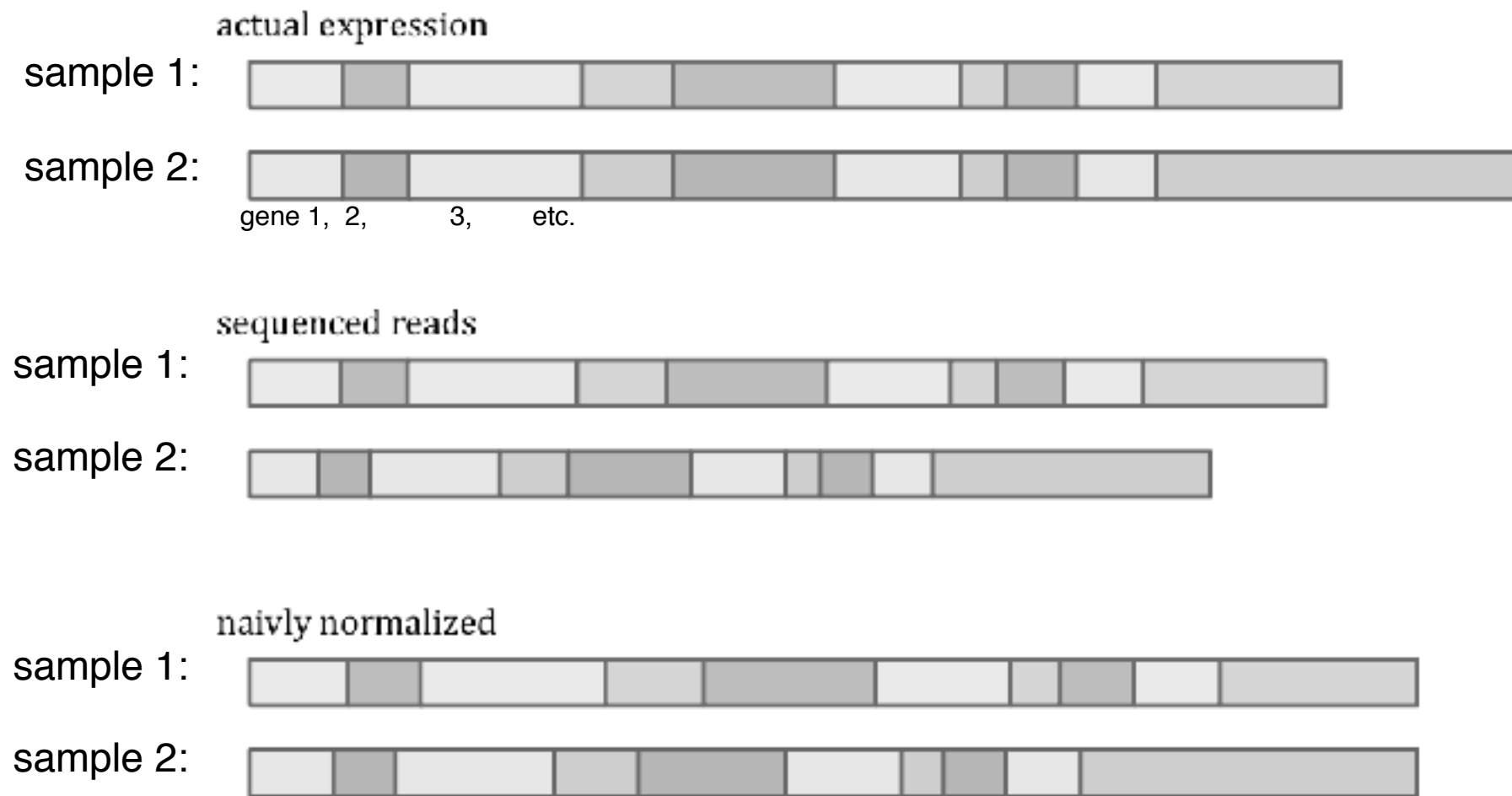
sample 1



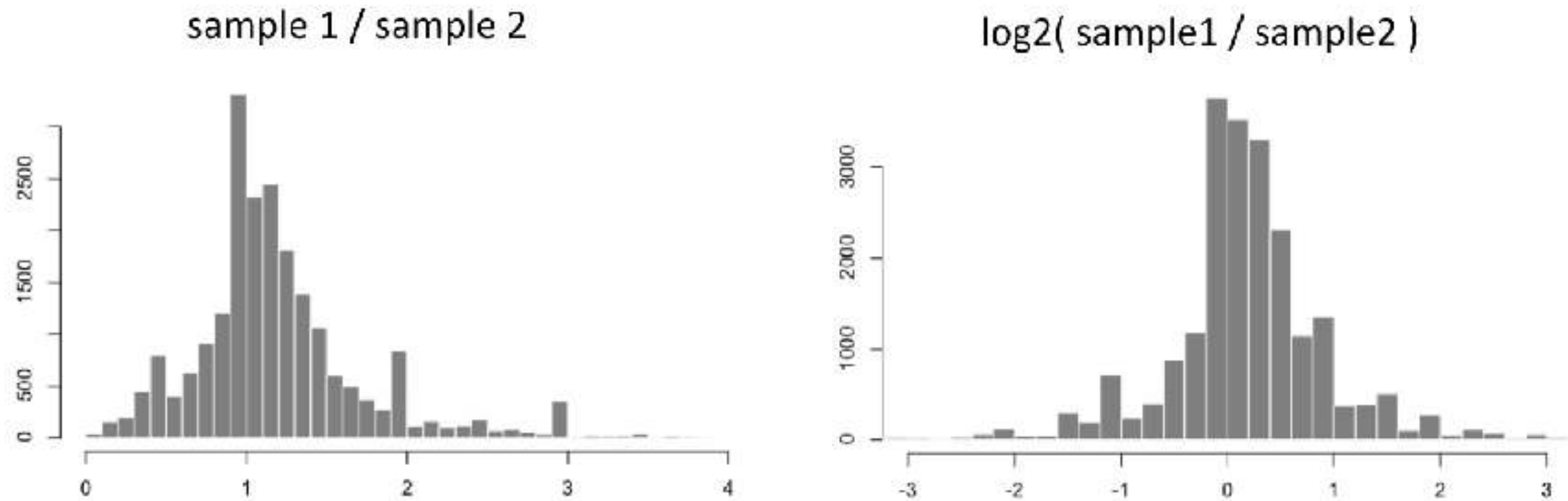
sample 2



Normalization for sequencing depth



DESeq2 uses a median of ratios method



- Create a reference sample by calculating the geometric mean across samples for each gene.
- For each sample, take the ratio with respect to the reference sample.
- Take the median across genes for each sample.

DESeq2 scaling factors or normalization factors?

$$S_j$$

"size factor"

per sample j

sequencing depth

robustly estimated
with median ratio

$$S_{ij}$$

"normalization factor"

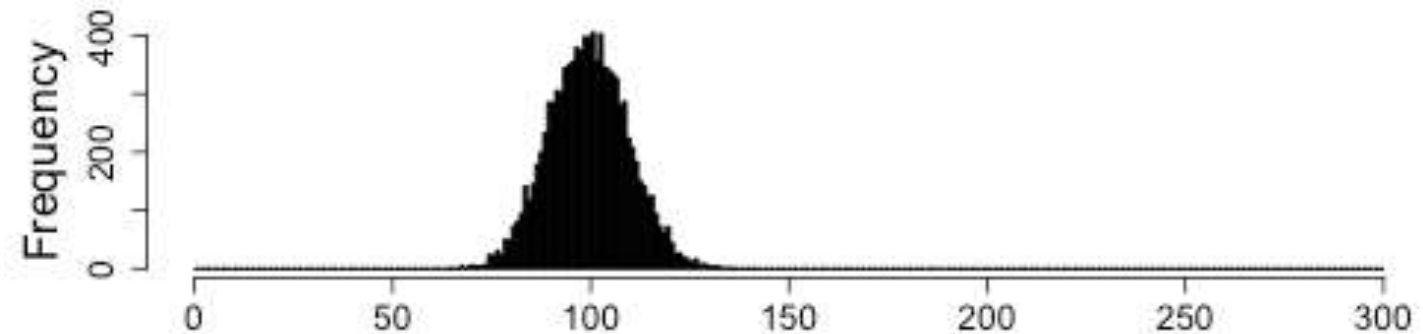
per sample j & per gene i

sequencing depth and
other factors differ across samples
(technical bias: [cqn](#) or [EDASeq](#))
(gene length: [tximport](#))

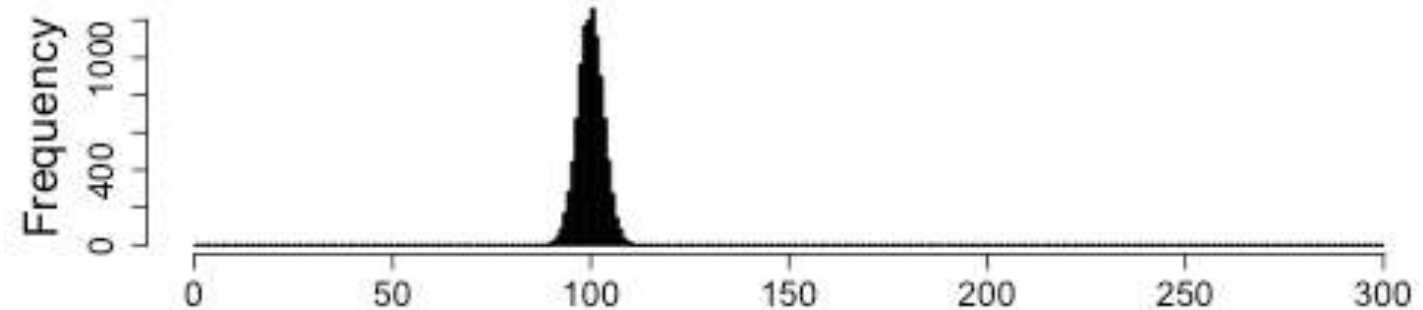
median ratio method for
sequencing depth can be
estimated on top

But, why counts and not other transformed data (e.g. FPKMs)?

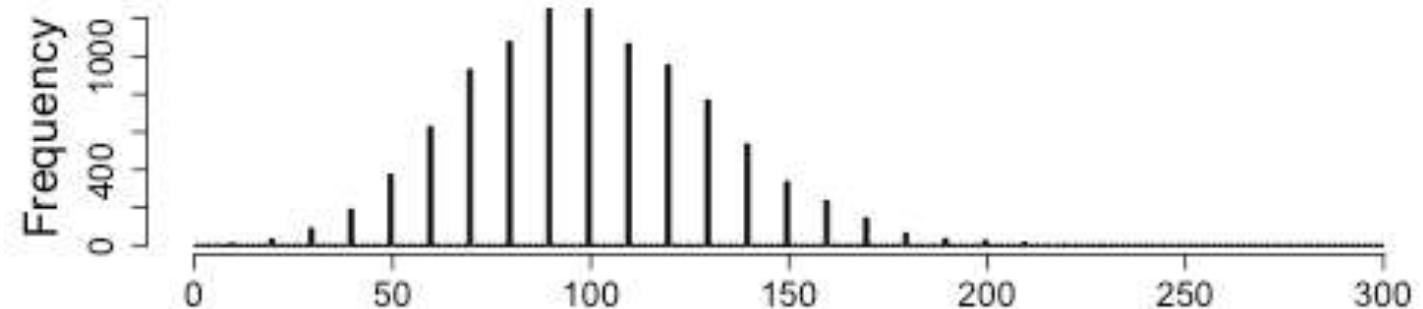
Raw count with mean of 100
Poisson sampling, so
SD=10

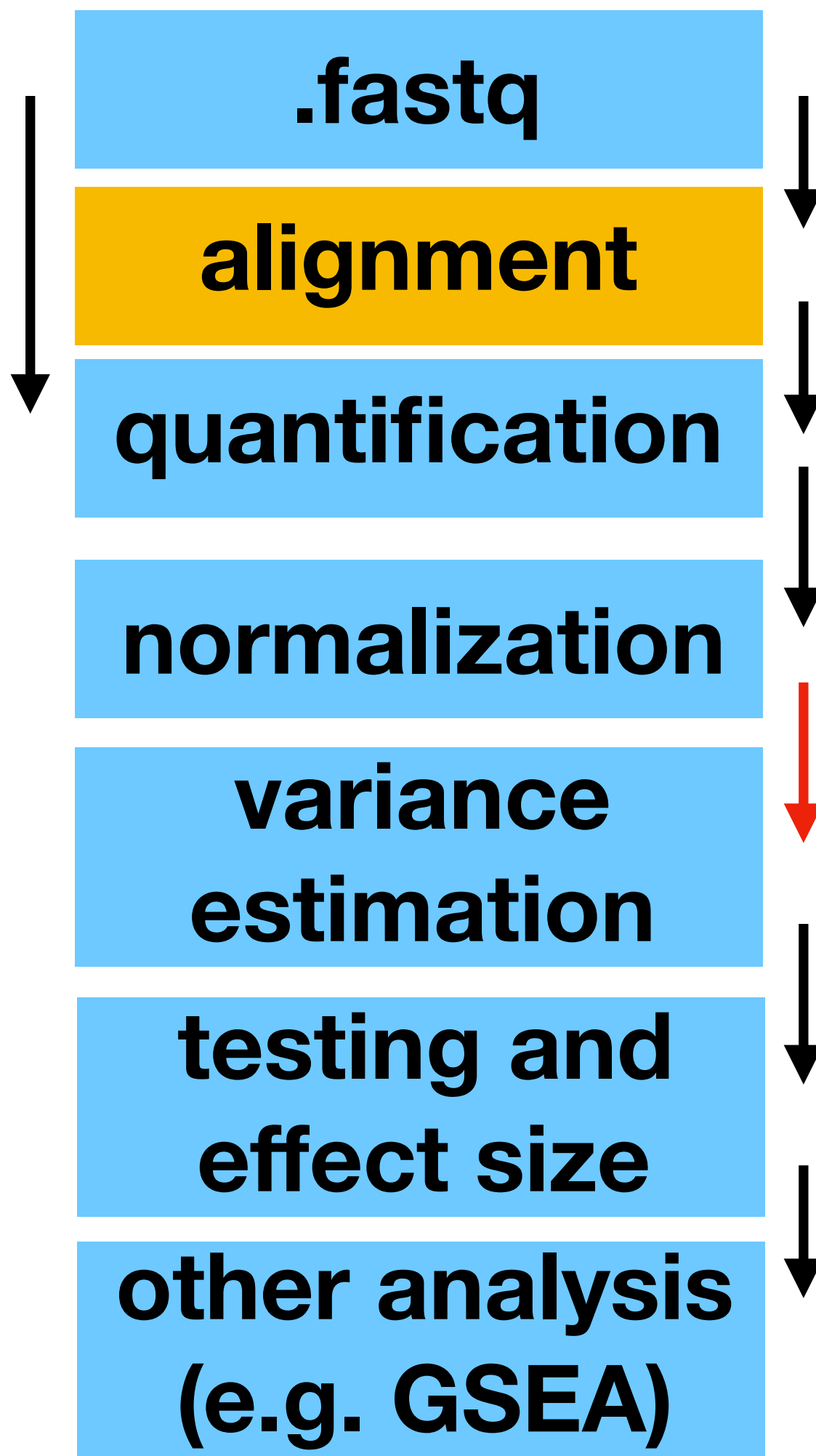


Raw count with mean of 100
scale by 1/10
SD = ?

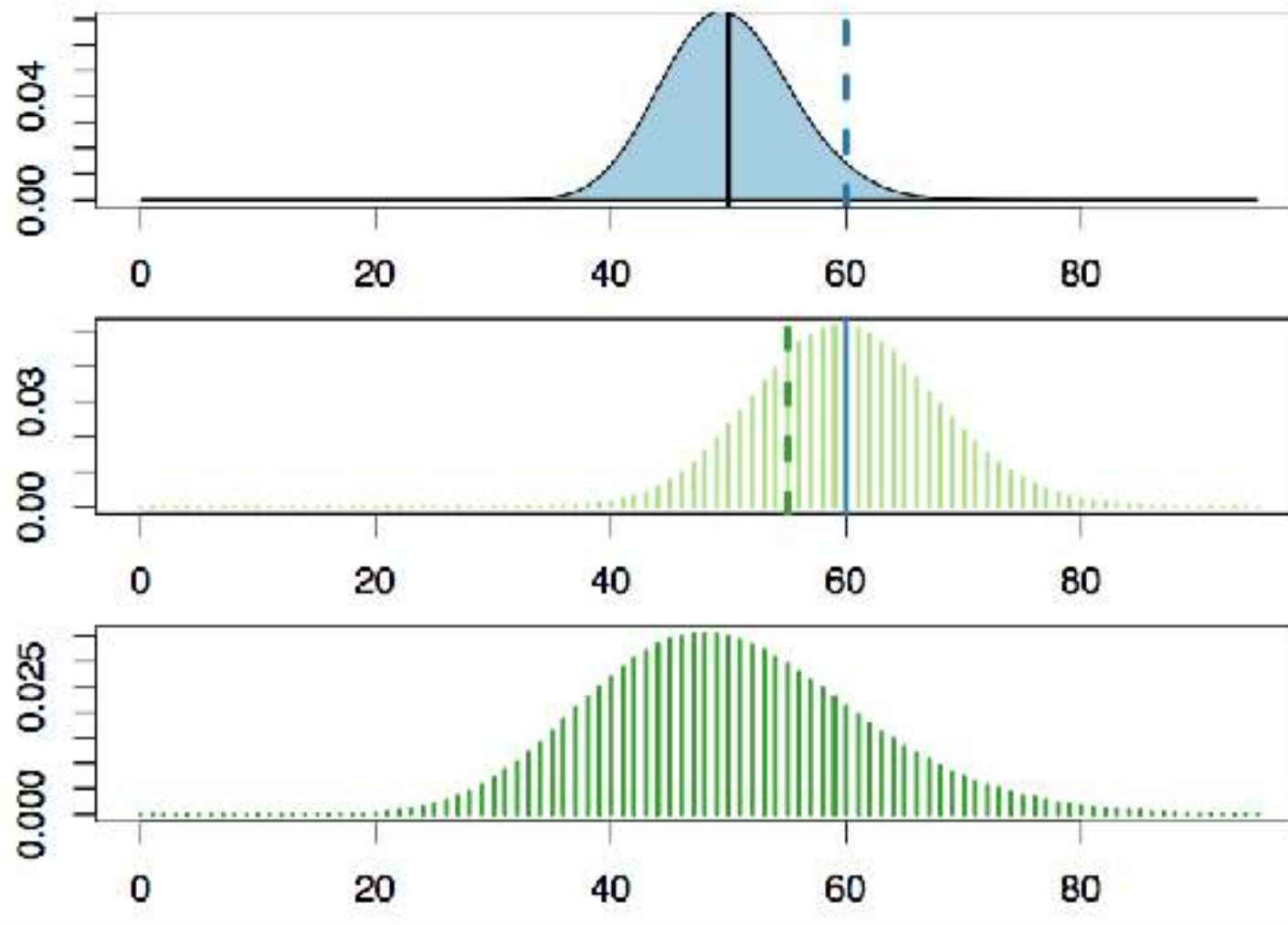


Raw count with mean of 100
scale by 10
SD = ?





Variance of a gene: technical noise + biological noise



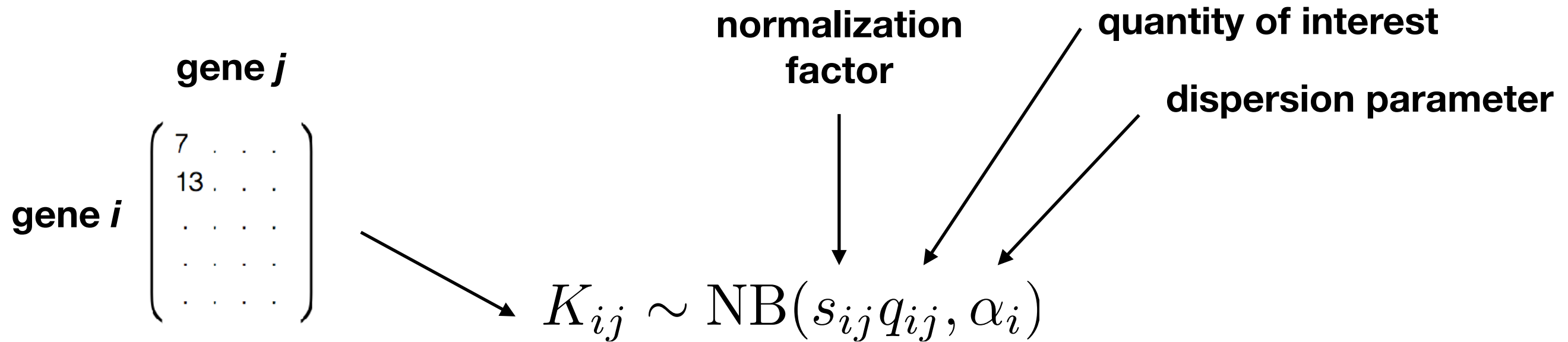
Biological sample with variance Γ

Poisson sampling variance = Λ

Gamma-poisson process

$$\text{NB}(\mu, \sigma^2 + \mu) = \Lambda(\Gamma(\mu, \sigma^2))$$

Poisson process in which the mean is gamma distributed



$$\text{Var}(K_{ij}) = \underbrace{\mu_{ij}}_{\text{technical noise}} + \alpha_i \underbrace{\mu_{ij}^2}_{\text{extra biological variability}}$$

technical noise

extra biological variability

The equation shows the variance of the count K_{ij} as the sum of two components. The first component, μ_{ij} , is labeled 'technical noise' with an arrow. The second component, $\alpha_i \mu_{ij}^2$, is labeled 'extra biological variability' with an arrow. Braces are used to group μ_{ij} and $\alpha_i \mu_{ij}^2$ under their respective labels.

Challenge: small number of replicates!

Dispersion estimation

$$\hat{\alpha}_{\text{MLE}} = \underset{\alpha}{\operatorname{argmax}}(\ell(\alpha|\vec{k}, \hat{\mu}))$$

$$\text{CR}(\alpha) = -\frac{1}{2} \log(\det(X^t W X))$$

$$\hat{\alpha}_{\text{CR}} = \underset{\alpha}{\operatorname{argmax}}(\ell(\alpha|\vec{k}, \hat{\mu}) + \text{CR}(\alpha))$$

$$\text{prior}(\alpha) = f_{\mathcal{N}}(\log(\alpha); \log(\alpha_{\text{fit}}), \sigma_{\alpha\text{-prior}}^2)$$

$$\hat{\alpha}_{\text{CR-MAP}} = \underset{\alpha}{\operatorname{argmax}}(\ell(\alpha|\vec{k}, \hat{\mu}) + \text{CR}(\alpha) + \log(\text{prior}(\alpha)))$$

Dispersion estimation

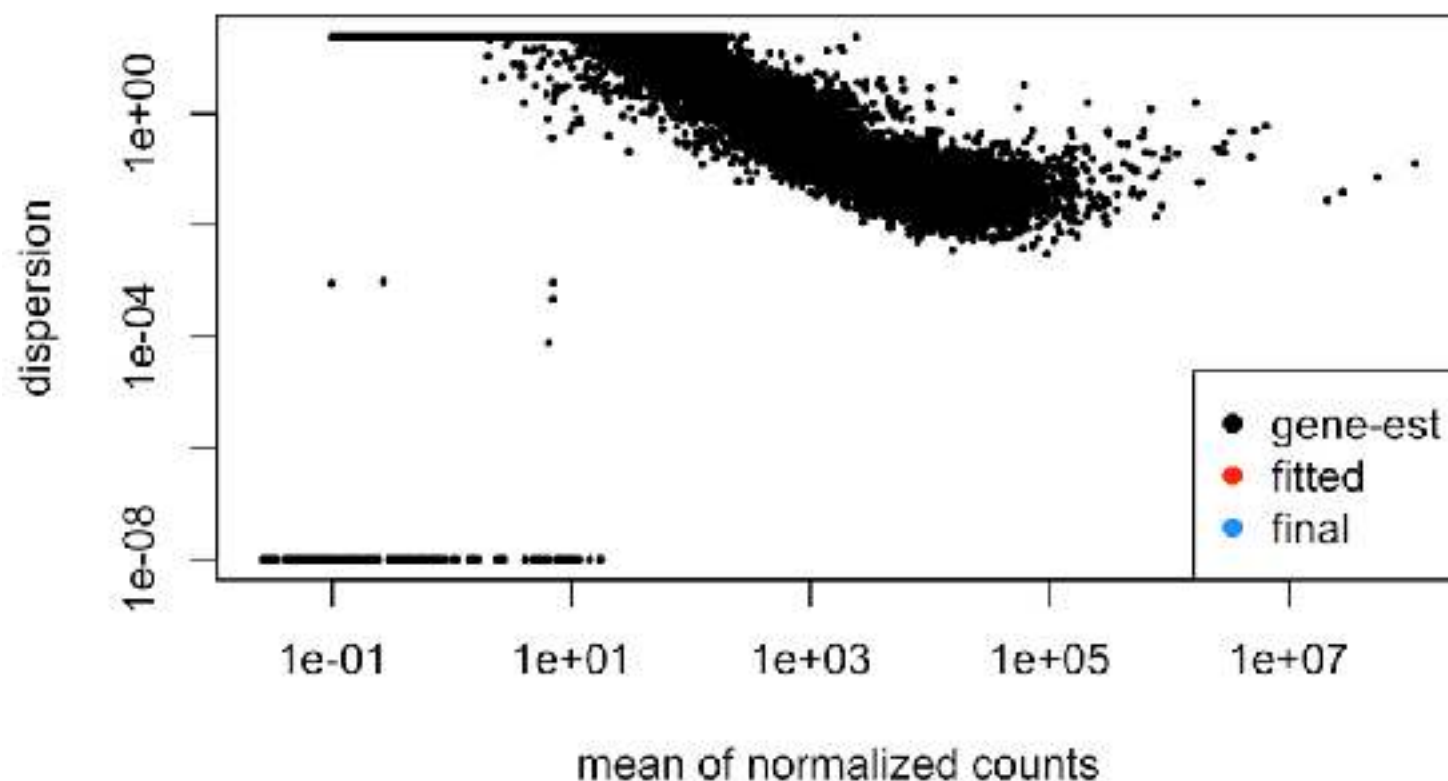
$$\hat{\alpha}_{\text{MLE}} = \underset{\alpha}{\operatorname{argmax}}(\ell(\alpha|\vec{k}, \hat{\mu})) \quad \longleftarrow \quad \text{maximum-likelihood estimates}$$

$$\text{CR}(\alpha) = -\frac{1}{2} \log(\det(X^t W X)) \quad \longleftarrow \quad \text{Cox-Reid bias term}$$

$$\hat{\alpha}_{\text{CR}} = \underset{\alpha}{\operatorname{argmax}}(\ell(\alpha|\vec{k}, \hat{\mu}) + \text{CR}(\alpha)) \quad \longleftarrow \quad \text{Cox-Reid ML estimate}$$

$$\text{prior}(\alpha) = f_{\mathcal{N}}(\log(\alpha); \log(\alpha_{\text{fit}}), \sigma_{\alpha\text{-prior}}^2)$$

$$\hat{\alpha}_{\text{CR-MAP}} = \underset{\alpha}{\operatorname{argmax}}(\ell(\alpha|\vec{k}, \hat{\mu}) + \text{CR}(\alpha) + \log(\text{prior}(\alpha)))$$



Dispersion estimation

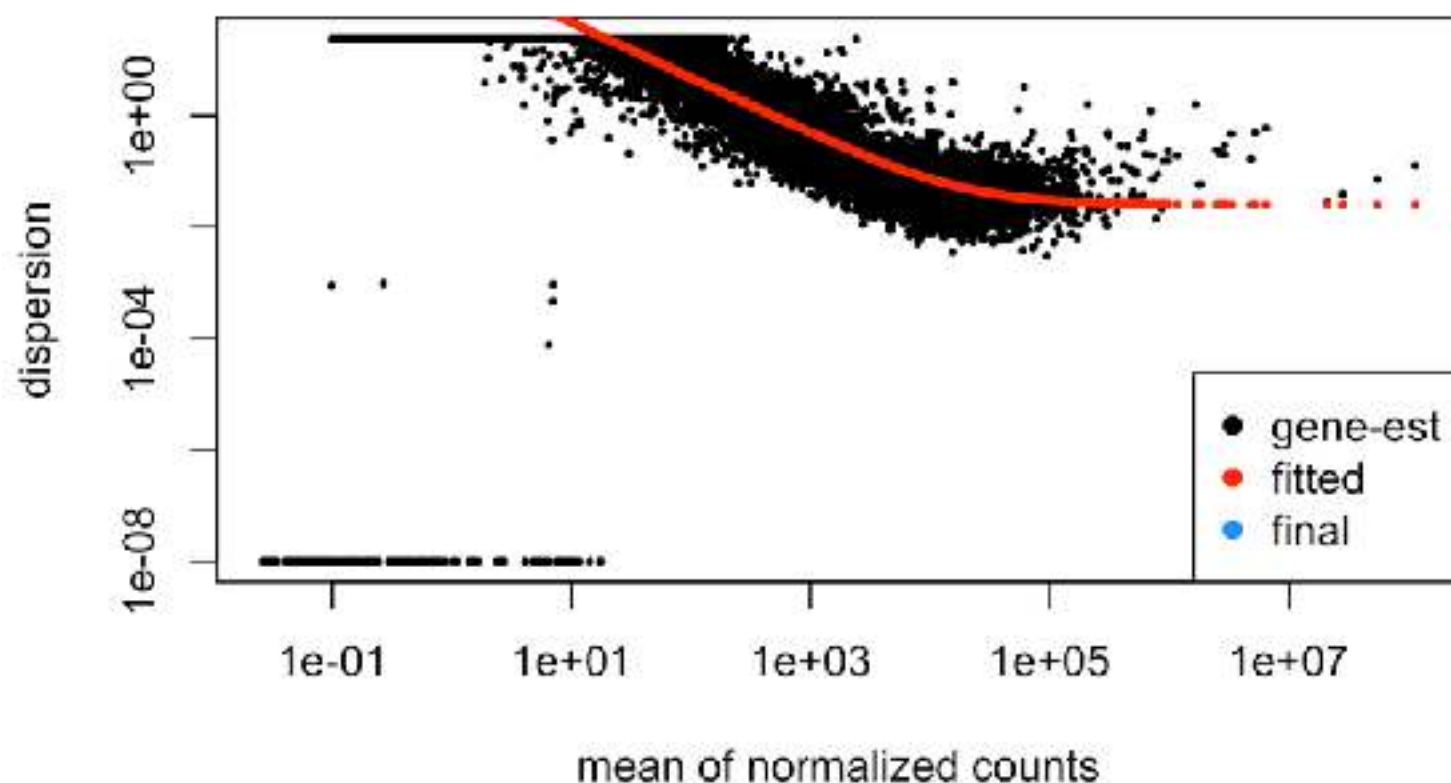
$$\hat{\alpha}_{\text{MLE}} = \underset{\alpha}{\operatorname{argmax}}(\ell(\alpha|\vec{k}, \hat{\mu})) \quad \longleftarrow \quad \text{maximum-likelihood estimates}$$

$$\text{CR}(\alpha) = -\frac{1}{2} \log(\det(X^t W X)) \quad \longleftarrow \quad \text{Cox-Reid bias term}$$

$$\hat{\alpha}_{\text{CR}} = \underset{\alpha}{\operatorname{argmax}}(\ell(\alpha|\vec{k}, \hat{\mu}) + \text{CR}(\alpha)) \quad \longleftarrow \quad \text{Cox-Reid ML estimate}$$

$$\text{prior}(\alpha) = f_{\mathcal{N}}(\log(\alpha); \log(\alpha_{\text{fit}}), \sigma_{\alpha\text{-prior}}^2) \quad \longleftarrow \quad \text{alpha prior by information sharing across genes}$$

$$\hat{\alpha}_{\text{CR-MAP}} = \underset{\alpha}{\operatorname{argmax}}(\ell(\alpha|\vec{k}, \hat{\mu}) + \text{CR}(\alpha) + \log(\text{prior}(\alpha)))$$



Dispersion estimation

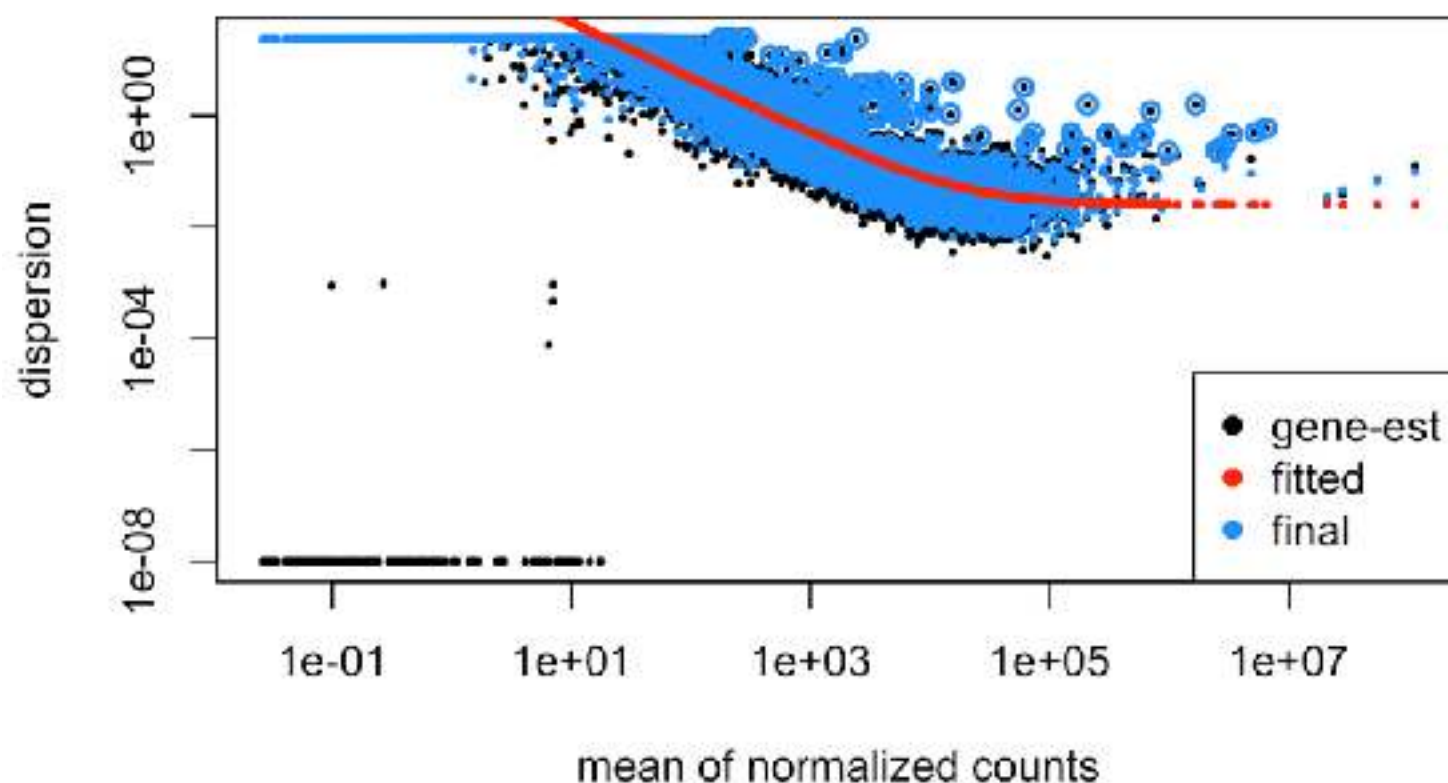
$$\hat{\alpha}_{\text{MLE}} = \underset{\alpha}{\operatorname{argmax}}(\ell(\alpha|\vec{k}, \hat{\mu})) \quad \leftarrow \text{maximum-likelihood estimates}$$

$$\text{CR}(\alpha) = -\frac{1}{2} \log(\det(X^t W X)) \quad \leftarrow \text{Cox-Reid bias term}$$

$$\hat{\alpha}_{\text{CR}} = \underset{\alpha}{\operatorname{argmax}}(\ell(\alpha|\vec{k}, \hat{\mu}) + \text{CR}(\alpha)) \quad \leftarrow \text{Cox-Reid ML estimate}$$

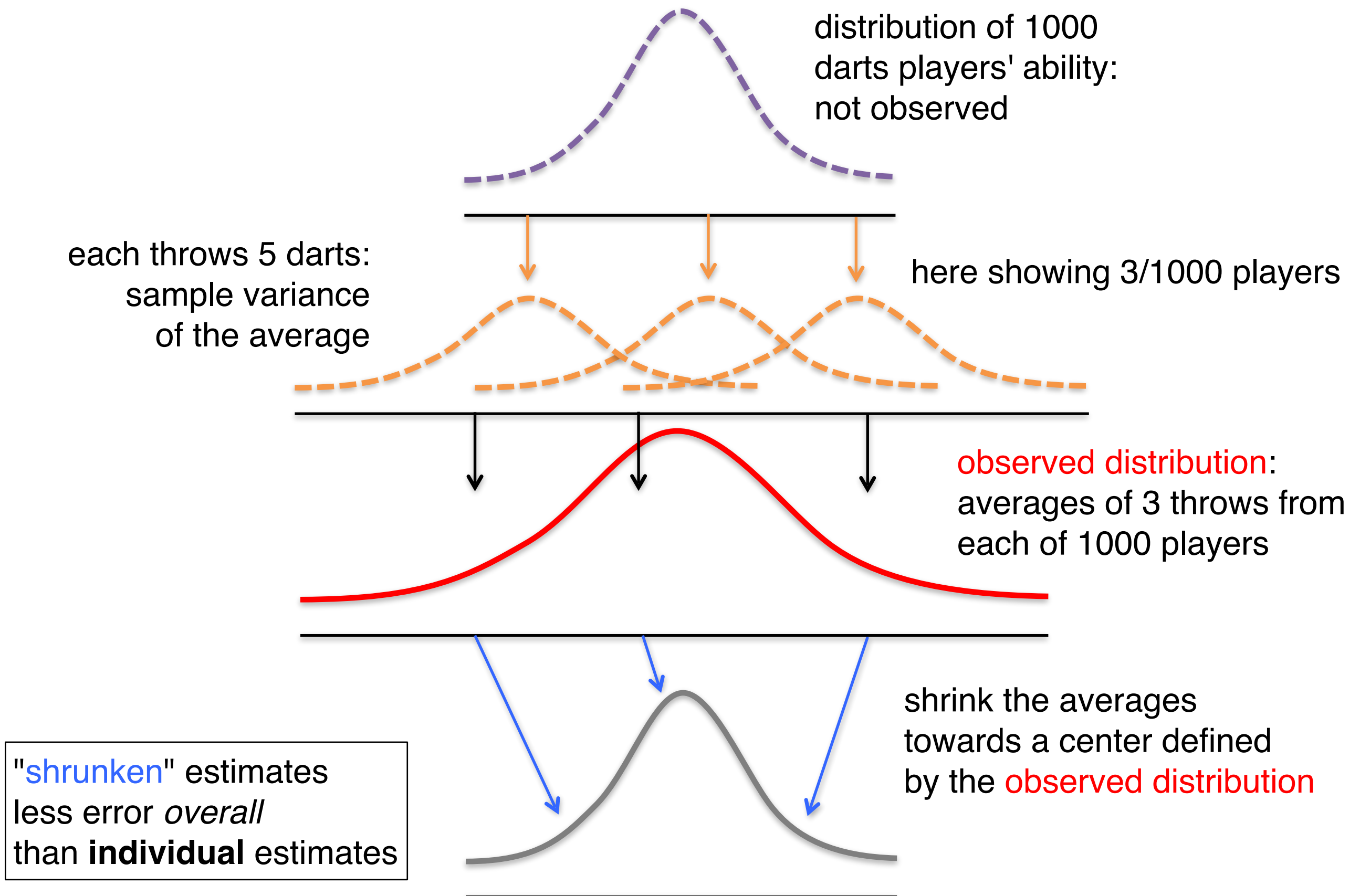
$$\text{prior}(\alpha) = f_{\mathcal{N}}(\log(\alpha); \log(\alpha_{\text{fit}}), \sigma_{\alpha\text{-prior}}^2) \quad \leftarrow \text{alpha prior by information sharing across genes}$$

$$\hat{\alpha}_{\text{CR-MAP}} = \underset{\alpha}{\operatorname{argmax}}(\ell(\alpha|\vec{k}, \hat{\mu}) + \text{CR}(\alpha) + \log(\text{prior}(\alpha)))$$



maximum a posteriori,
penalized likelihood

Why shrinkage?



Shrinkage estimation

population
distribution

dashed = unobserved

sampling variance
around true ability

empirical
distribution

the center defines
the prior mean

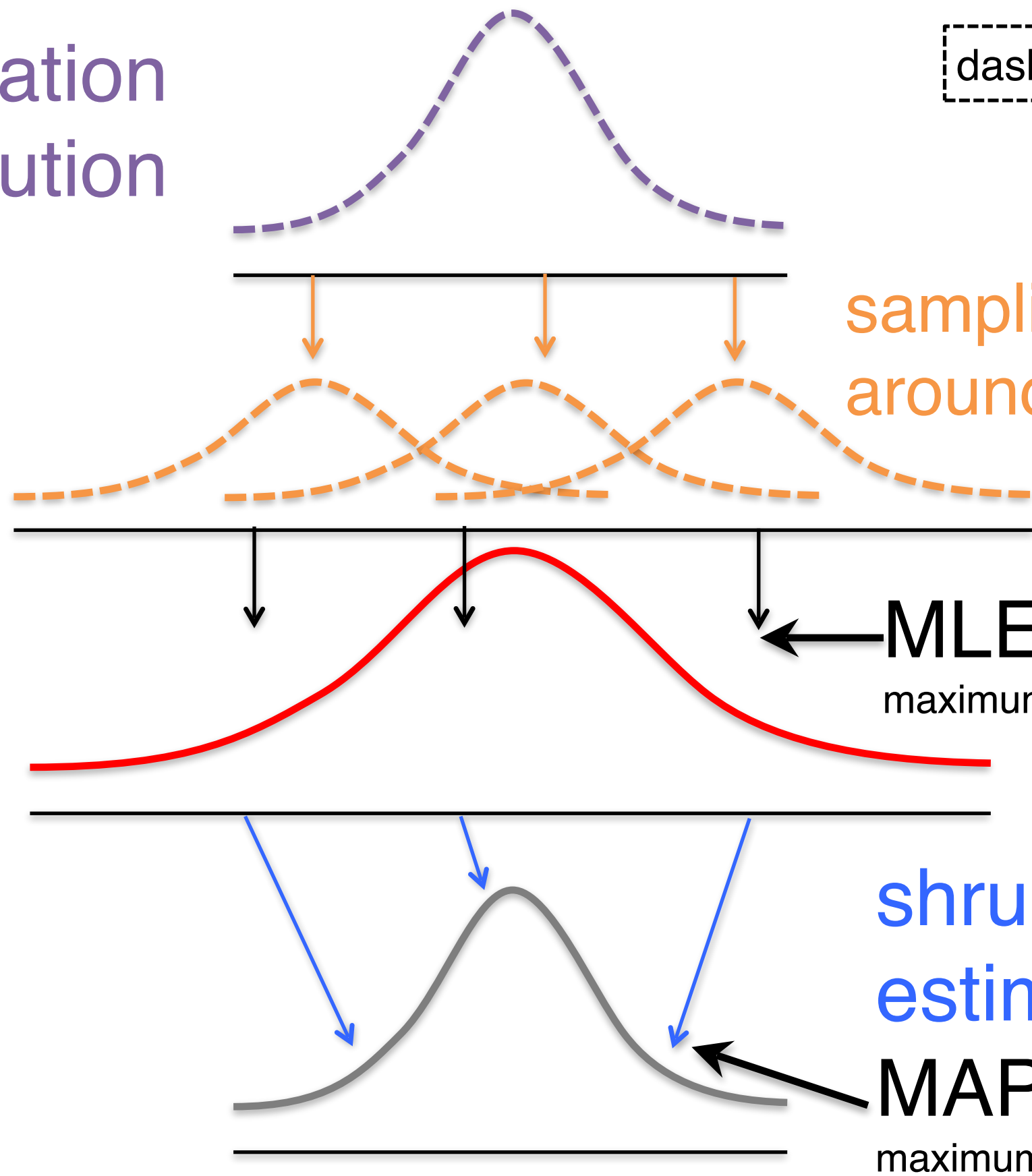
MLE

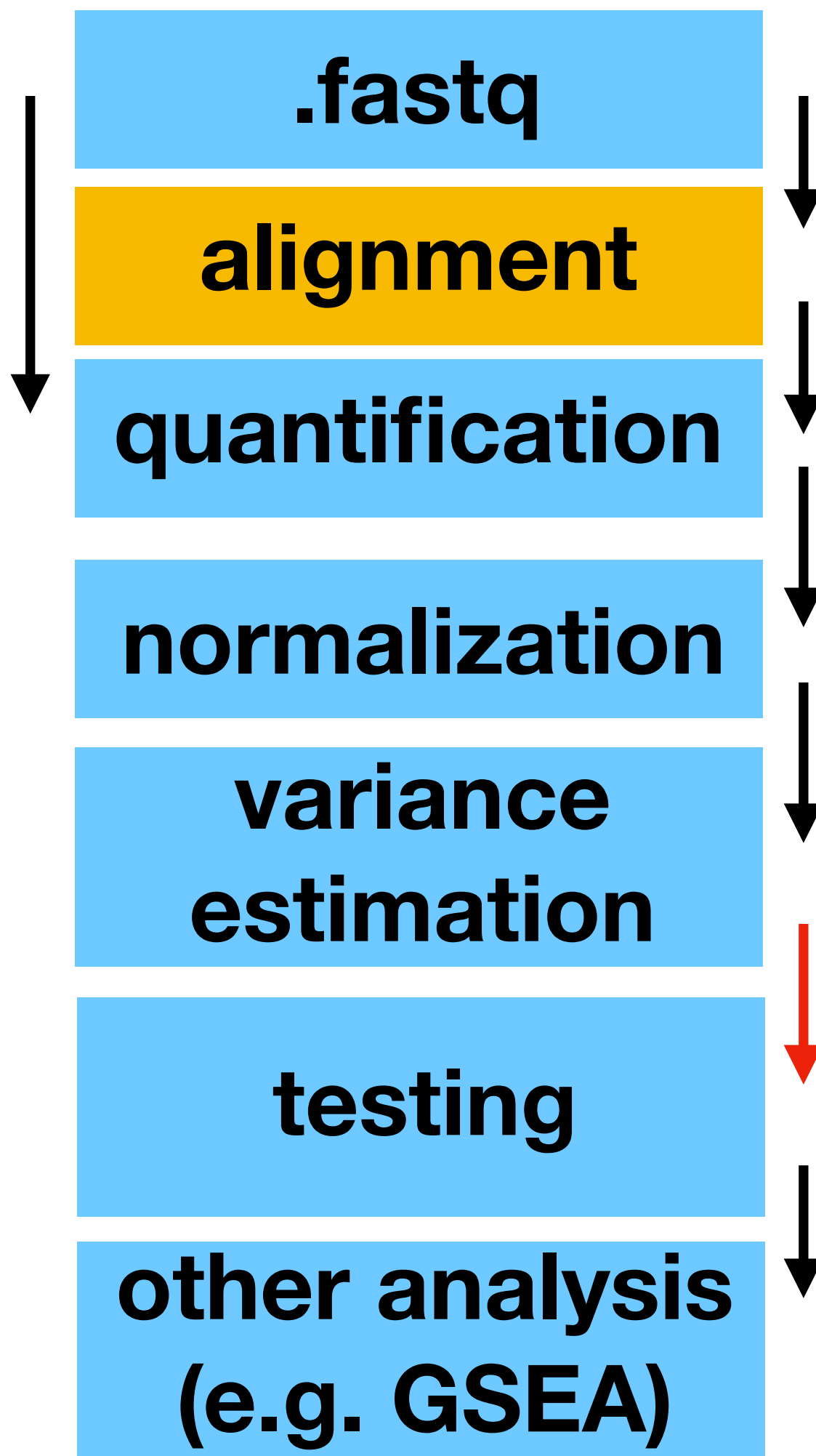
maximum likelihood estimates

shrunk
estimates or

MAP

maximum a posteriori

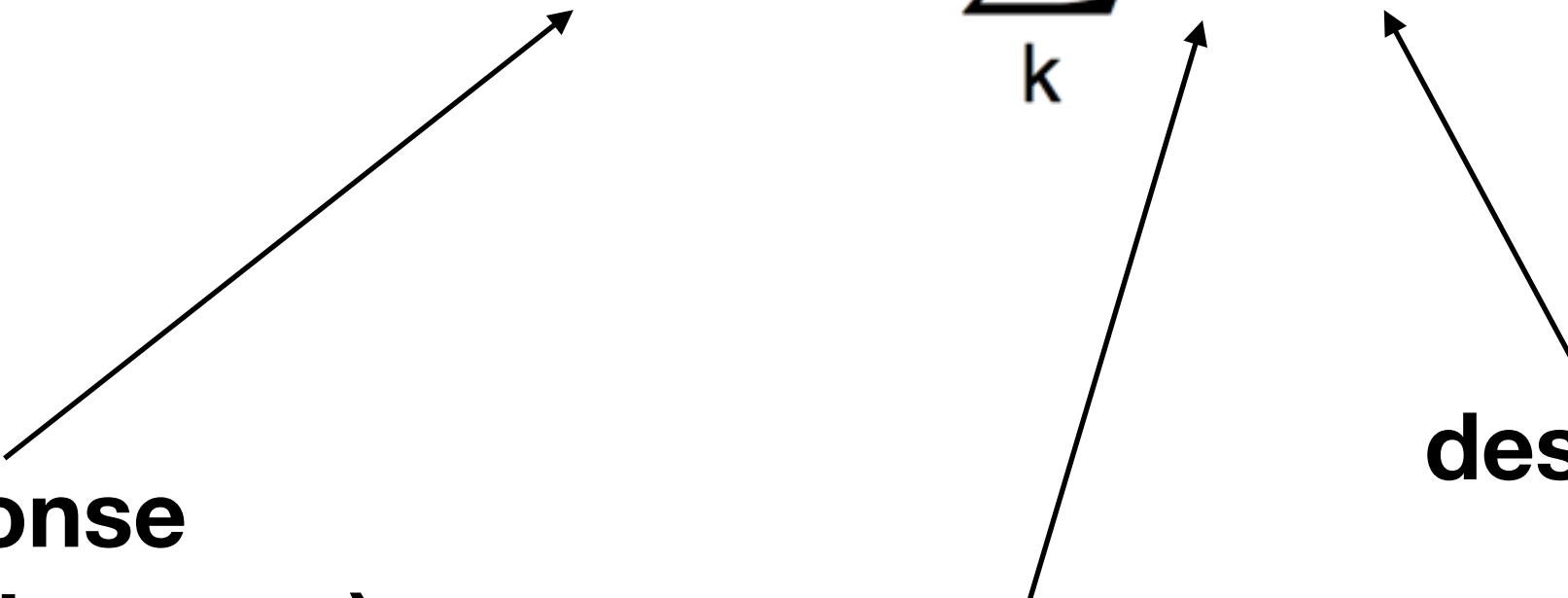




Generalized linear models

$$\log \mu_{ij} = \sum_k \beta_{ik} x_{kj}$$

response
(expected counts)



The diagram consists of three arrows pointing upwards towards the equation. The first arrow originates from the text 'response (expected counts)' and points to the term $\log \mu_{ij}$. The second arrow originates from the text 'predictors (or differential expression effects)' and points to the coefficient β_{ik} . The third arrow originates from the text 'design matrix' and points to the variable x_{kj} .

predictors
(or differential expression effects)

design matrix

Simplest case: two-group comparison

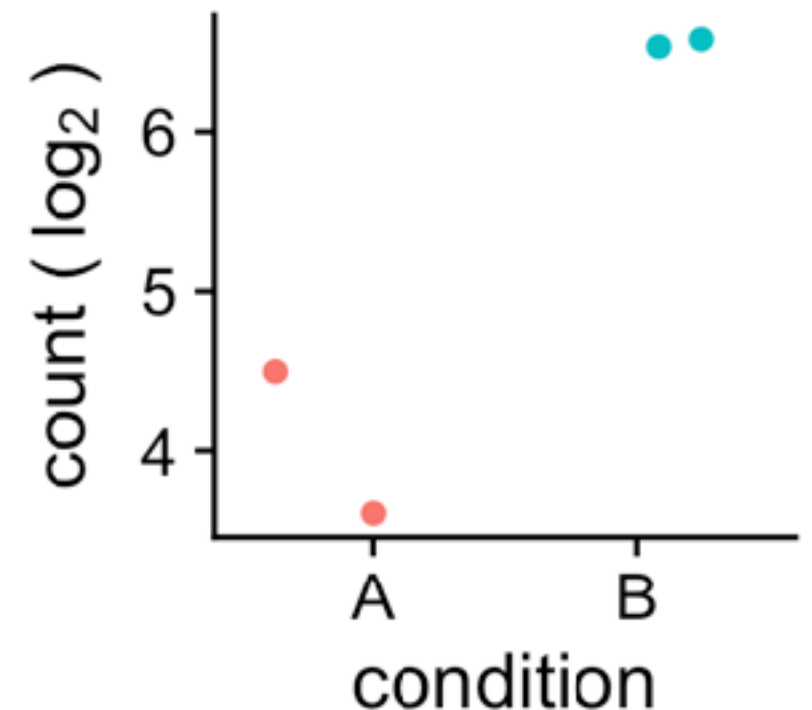
DESeq2 design = ~ condition

$$\begin{array}{cc} \text{condition} & \\ \mathbf{A} & \log_2 u_1 \\ \mathbf{A} & \log_2 u_2 \\ \mathbf{B} & \log_2 u_3 \\ \mathbf{B} & \log_2 u_4 \end{array} = \begin{array}{c} x_1 \\ \left(\begin{array}{cc} 1 & 0 \\ 1 & 0 \\ 1 & 1 \\ 1 & 1 \end{array} \right) \end{array} \begin{array}{c} \left(\begin{array}{c} \beta_{\text{Intercept}} \\ \beta_{\text{condition_A_vs_B}} \end{array} \right) \end{array}$$

estimated log expression

$$\text{condition A} = \beta_{\text{Intercept}}$$

$$\text{condition B} = \beta_{\text{Intercept}} + \beta_{\text{condition_A_vs_B}}$$



Multi-level comparisons

DESeq2 design = ~ condition

condition		x_1	x_2	
A	$\log_2 u_1$	1	0	$\begin{bmatrix} \beta_{\text{Intercept}} \\ \beta_{\text{condition_A_vs_B}} \\ \beta_{\text{condition_A_vs_C}} \end{bmatrix}$
A	$\log_2 u_2$	1	0	
B	$\log_2 u_3$	1	1	
B	$\log_2 u_4$	1	1	
C	$\log_2 u_5$	1	0	
C	$\log_2 u_6$	1	0	

=

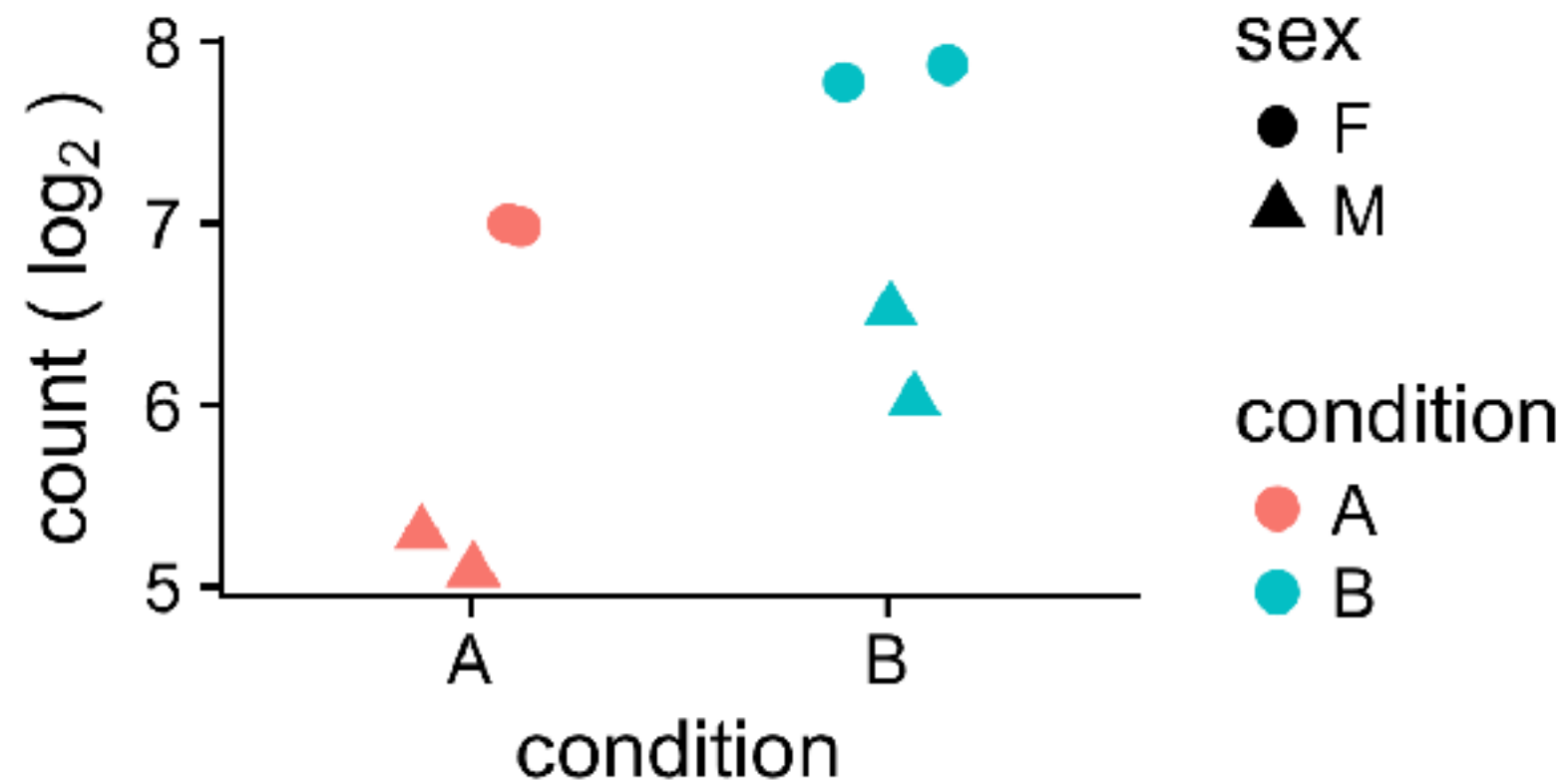
estimated log expression

condition A = $\beta_{\text{Intercept}}$

condition B = $\beta_{\text{Intercept}}$ + $\beta_{\text{condition_A_vs_B}}$

condition C = $\beta_{\text{Intercept}}$ + $\beta_{\text{condition_A_vs_C}}$

Comparisons with blocking factors



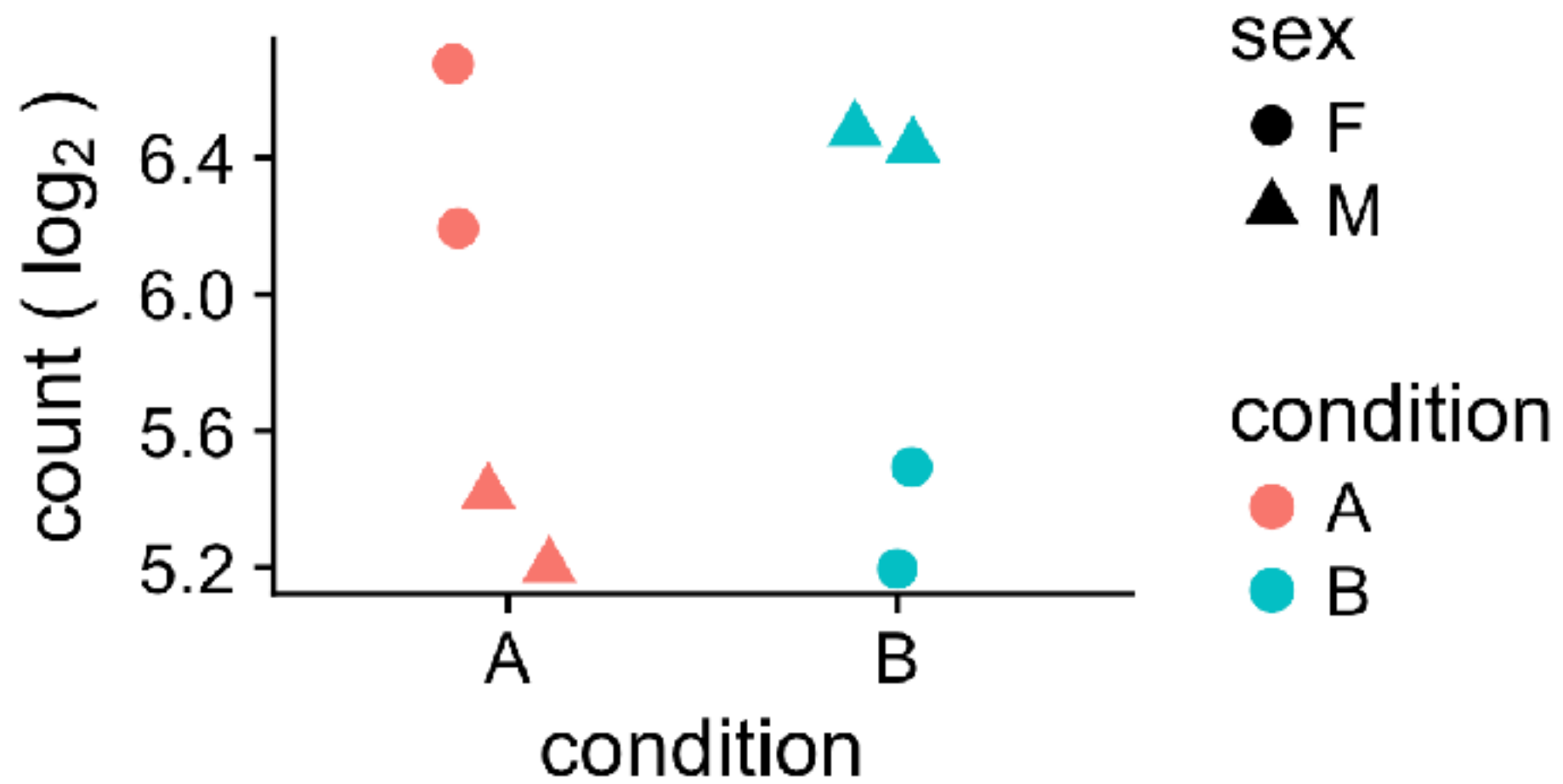
Comparisons with blocking factors

DESeq2 design = ~ sex + condition

sex	condition			
M	A	$\log_2 u_1$	=	$\begin{matrix} & x_1 & x_2 \\ \begin{pmatrix} 1 & 1 & 0 \\ 1 & 1 & 0 \\ 1 & 1 & 1 \\ 1 & 1 & 1 \\ 1 & 0 & 0 \\ 1 & 0 & 0 \\ 1 & 0 & 1 \\ 1 & 0 & 1 \end{pmatrix} & \begin{pmatrix} \beta_{\text{Intercept}} \\ \beta_{\text{sex_female_vs_male}} \\ \beta_{\text{condition_A_vs_B}} \end{pmatrix} \end{matrix}$
M	A	$\log_2 u_2$		
M	B	$\log_2 u_3$		
M	B	$\log_2 u_4$		
F	A	$\log_2 u_5$		
F	A	$\log_2 u_6$		
F	B	$\log_2 u_7$		
F	B	$\log_2 u_8$		

What would be the predicted log expression for females in condition B?

Interactions



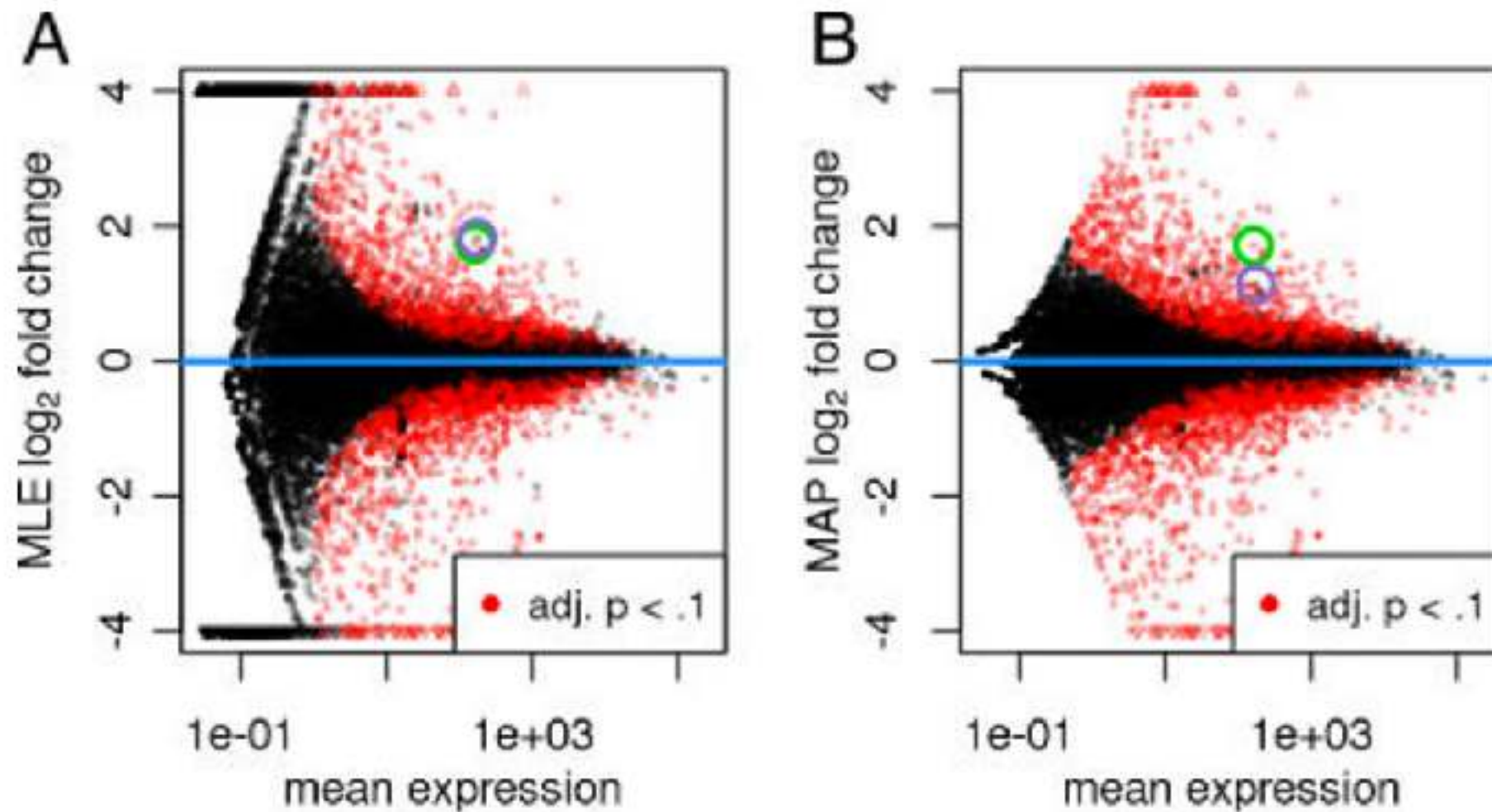
Comparisons with blocking factors

DESeq2 design = ~ sex + condition + sex:condition

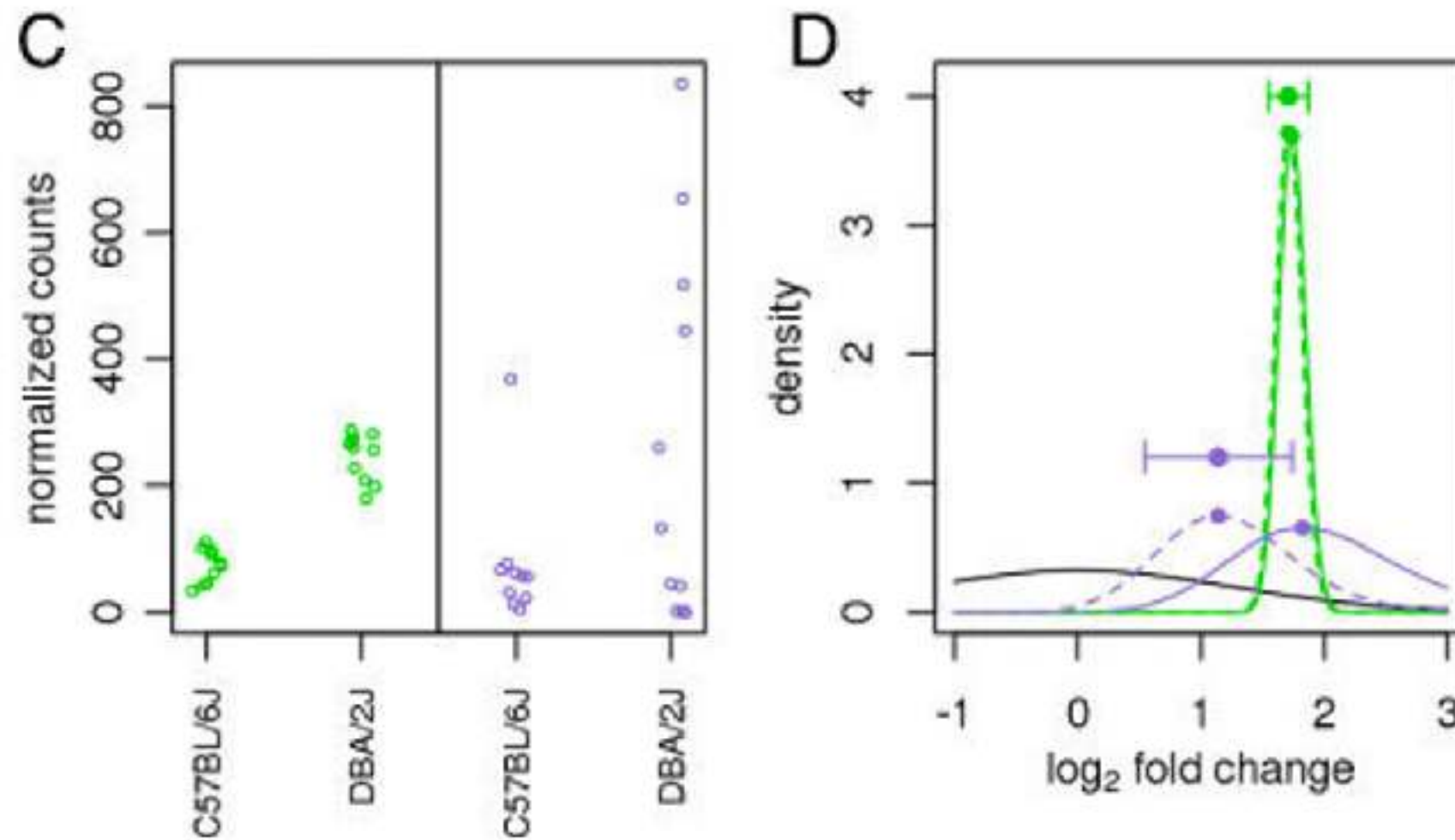
sex	condition	
M	A	$\log_2 u_1$
M	A	$\log_2 u_2$
M	B	$\log_2 u_3$
M	B	$\log_2 u_4$
F	A	$\log_2 u_5$
F	A	$\log_2 u_6$
F	B	$\log_2 u_7$
F	B	$\log_2 u_8$

$$= \begin{bmatrix} & X_1 & X_2 & X_3 \\ 1 & 1 & 0 & 0 \\ 1 & 1 & 0 & 0 \\ 1 & 1 & 1 & 1 \\ 1 & 1 & 1 & 1 \\ 1 & 0 & 0 & 0 \\ 1 & 0 & 0 & 0 \\ 1 & 0 & 1 & 0 \\ 1 & 0 & 1 & 0 \end{bmatrix} \begin{bmatrix} \beta_{\text{Intercept}} \\ \beta_{\text{sex_female_vs_male}} \\ \beta_{\text{condition_A_vs_B}} \\ \beta_{\text{interaction}} \end{bmatrix}$$

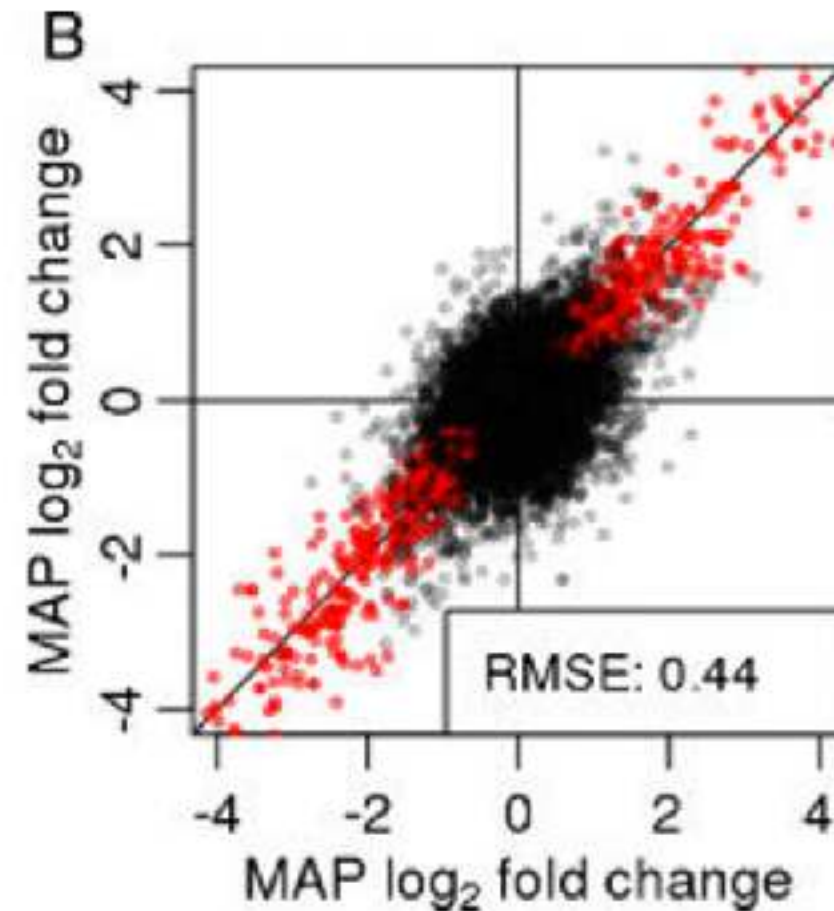
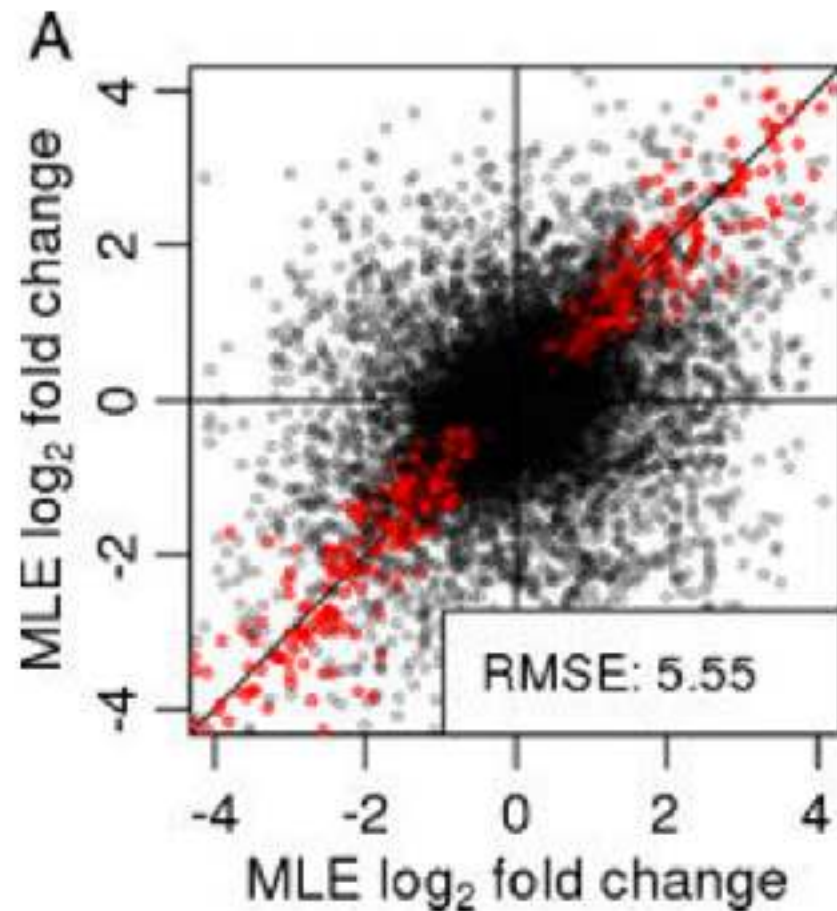
DESeq2 shrinkage of log fold changes



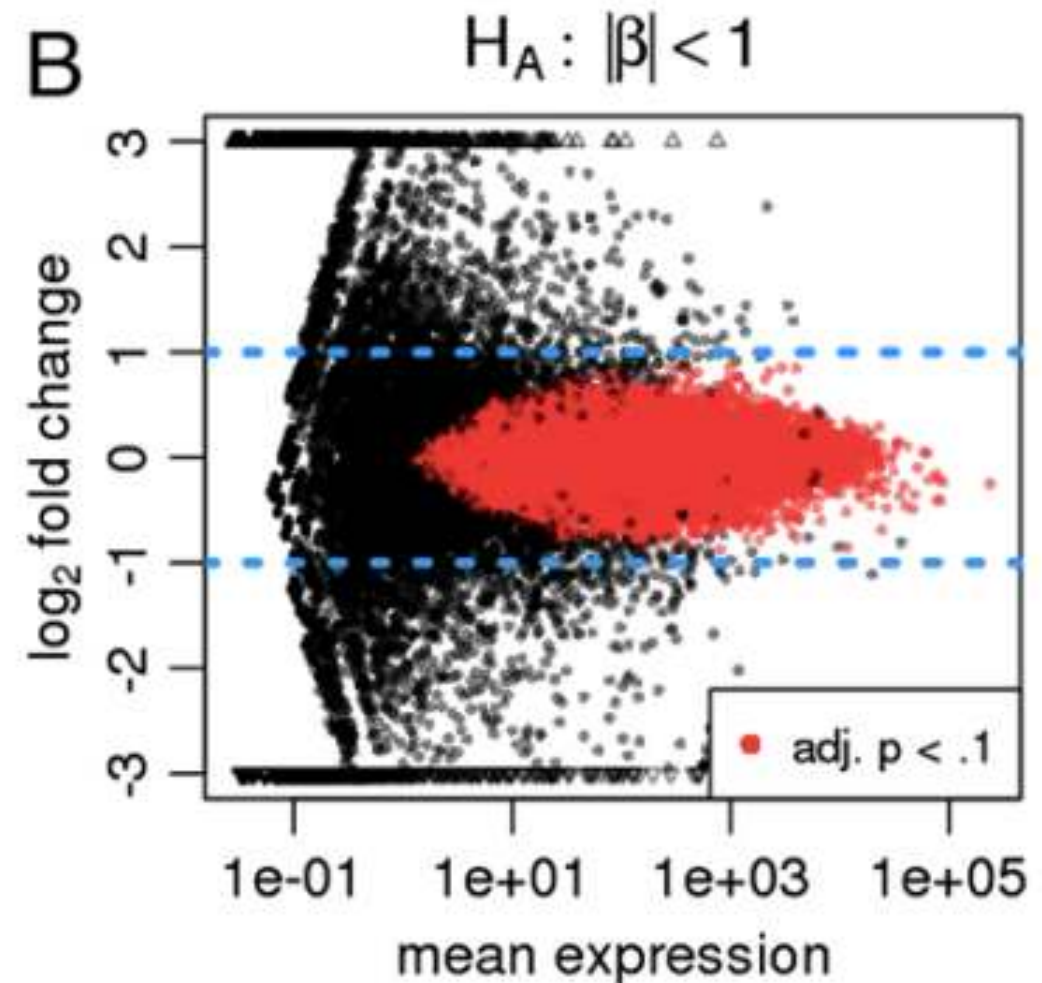
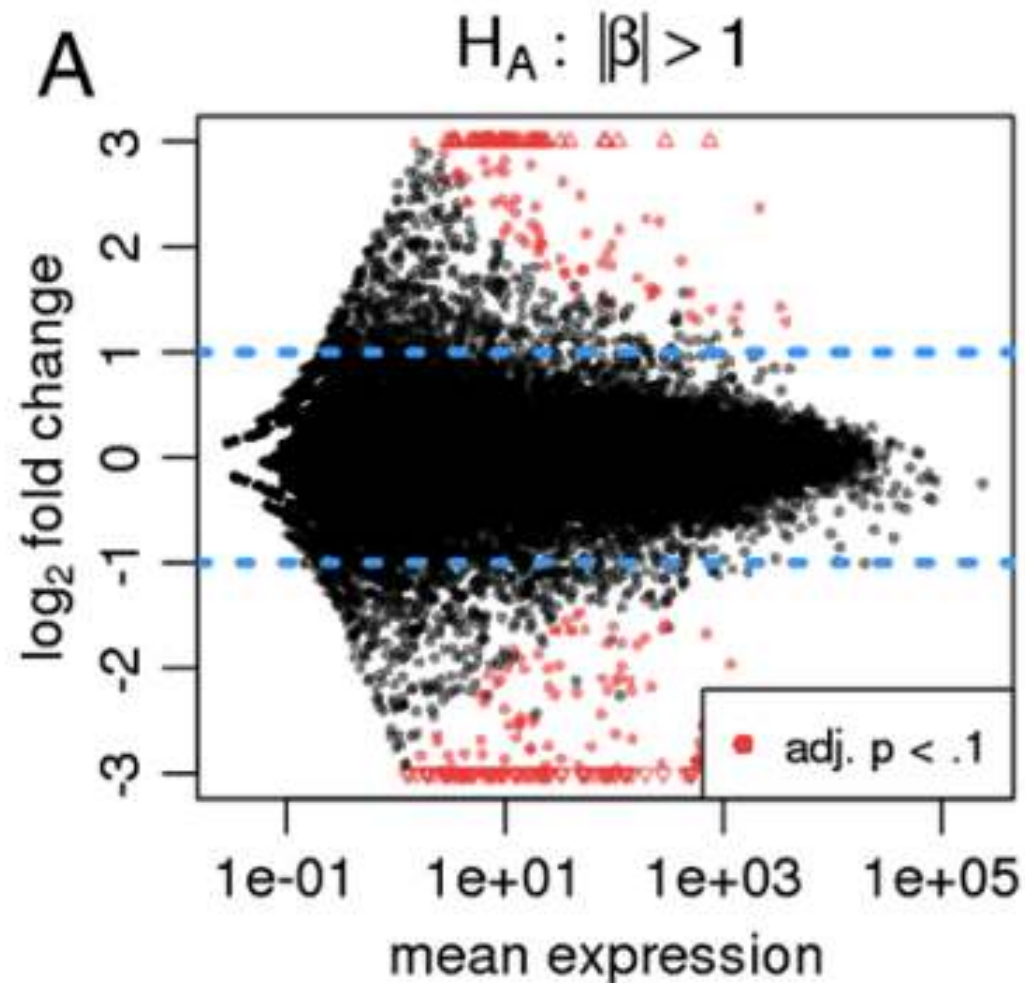
DESeq2 shrinkage of log fold changes



DESeq2 shrinkage of log fold changes



Other hypothesis testing options



Confounders and experimental design

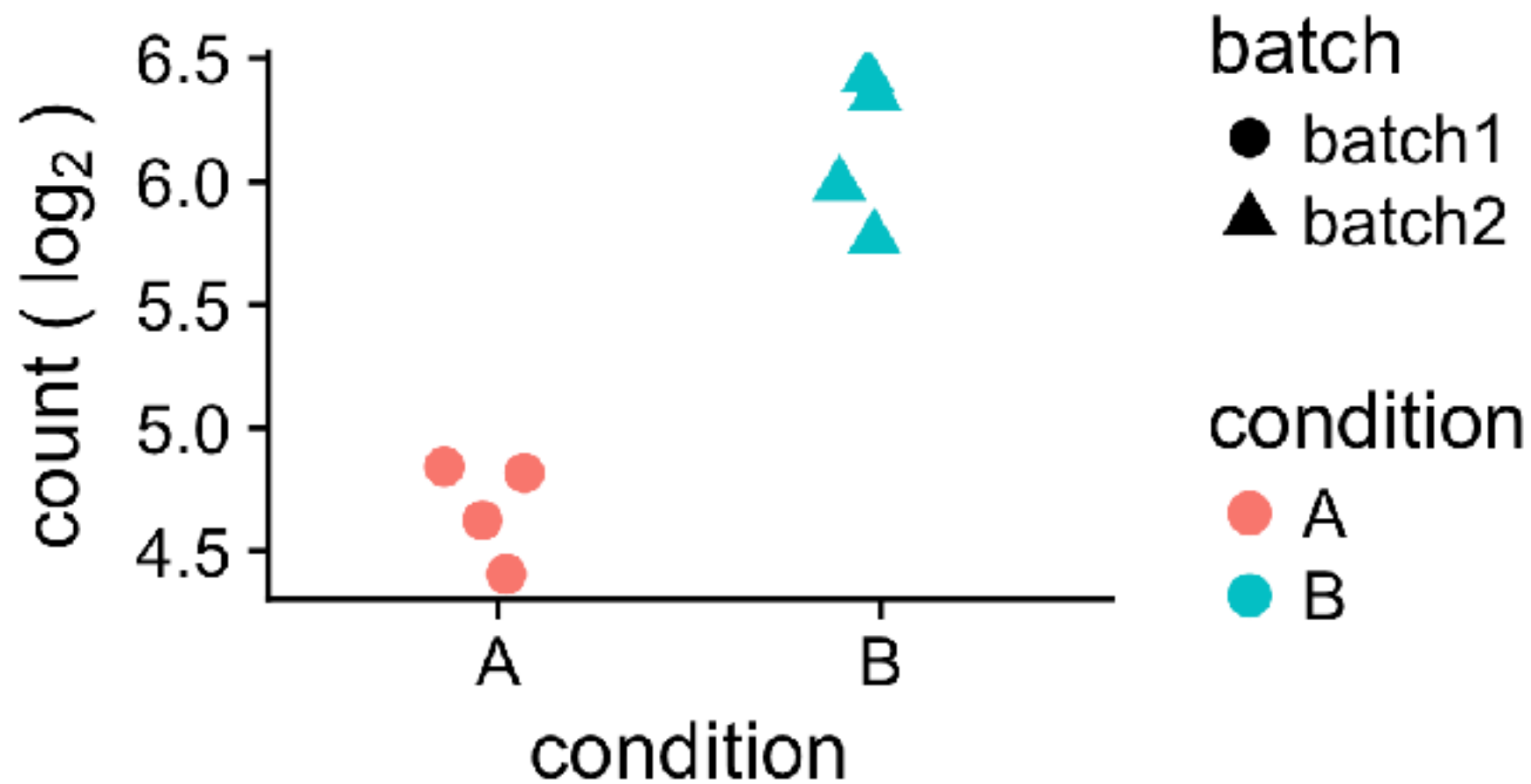
- Spielman et al., Nature Genetics 2007

78% of genes are differentially expressed between human populations

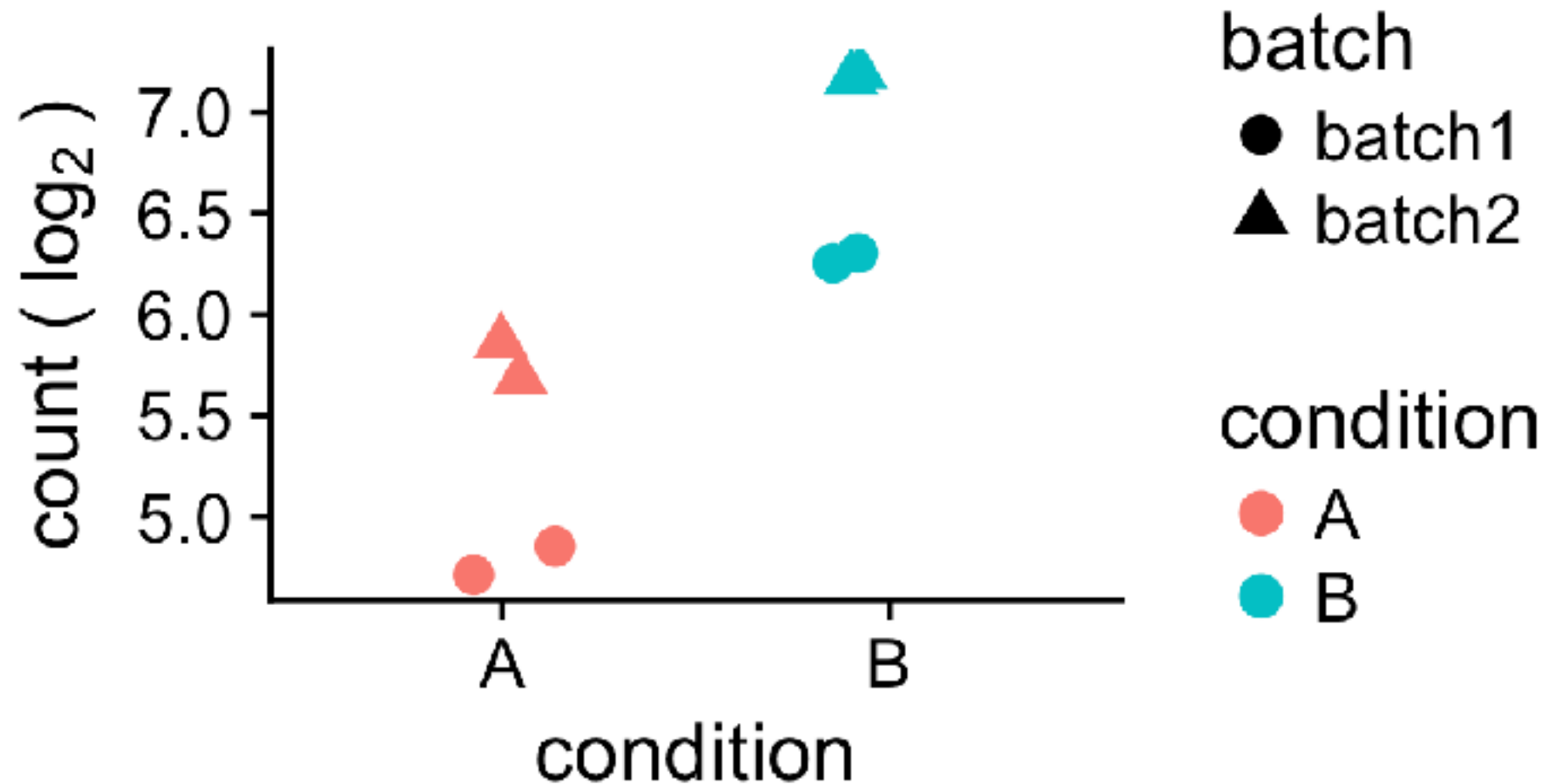
- Lin et al., PNAS, 2014

Differences in gene expression across species are larger than between tissues of a same species.

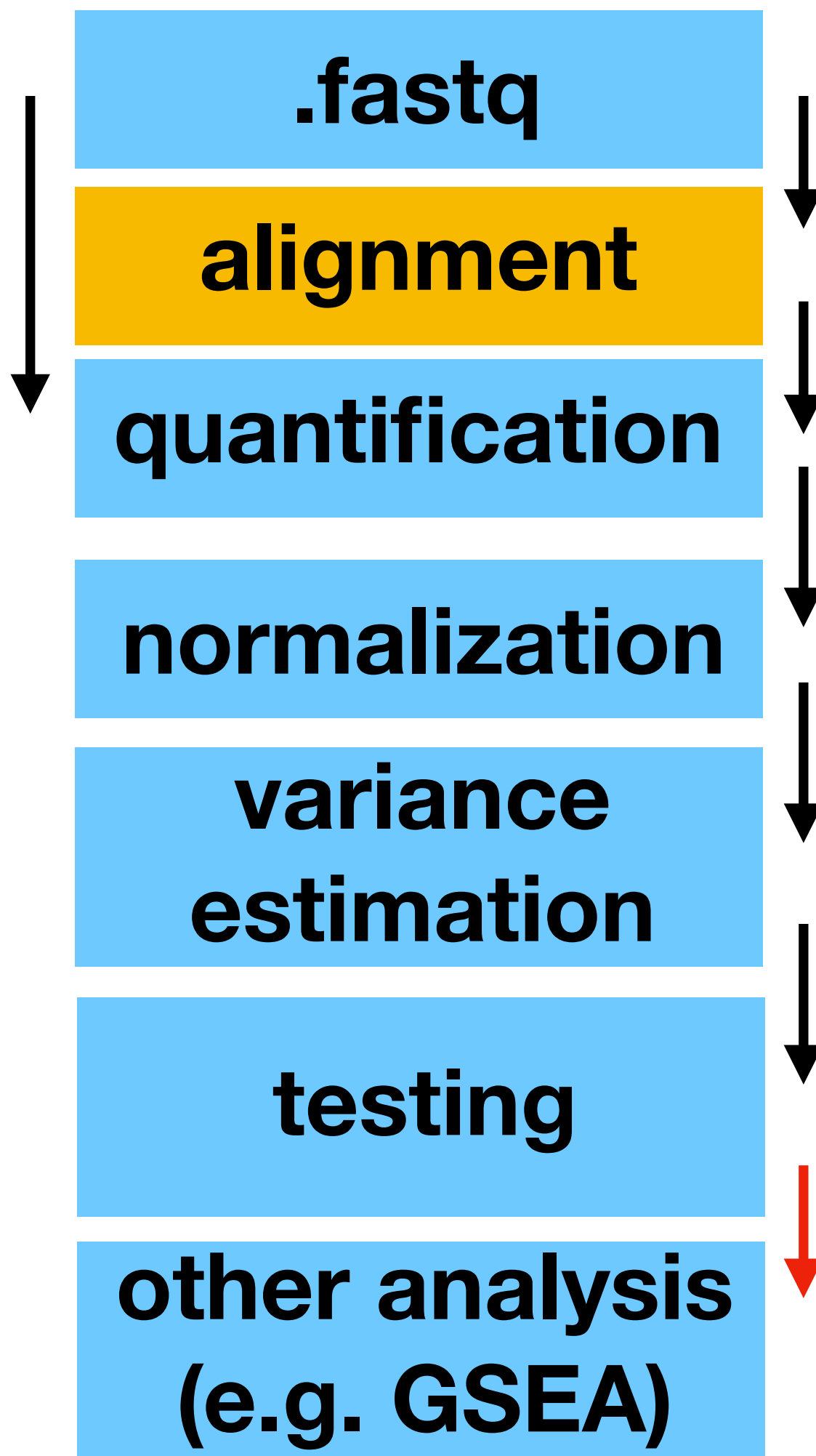
Confounders experiment



Confounders and experimental design



- Randomize conditions of interest in batches and include it as a blocking factor.



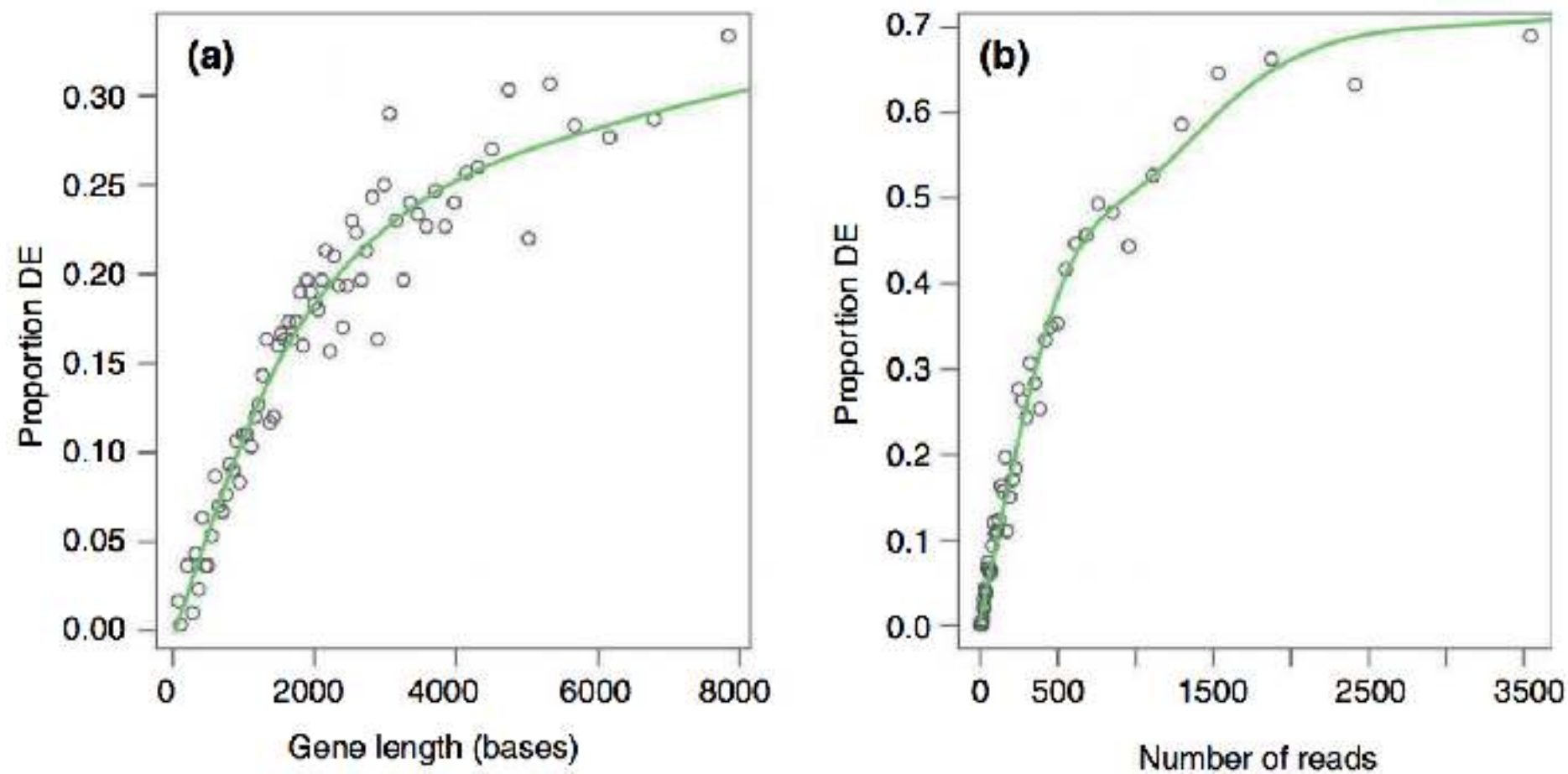
“Traditional” gene-set enrichment analysis

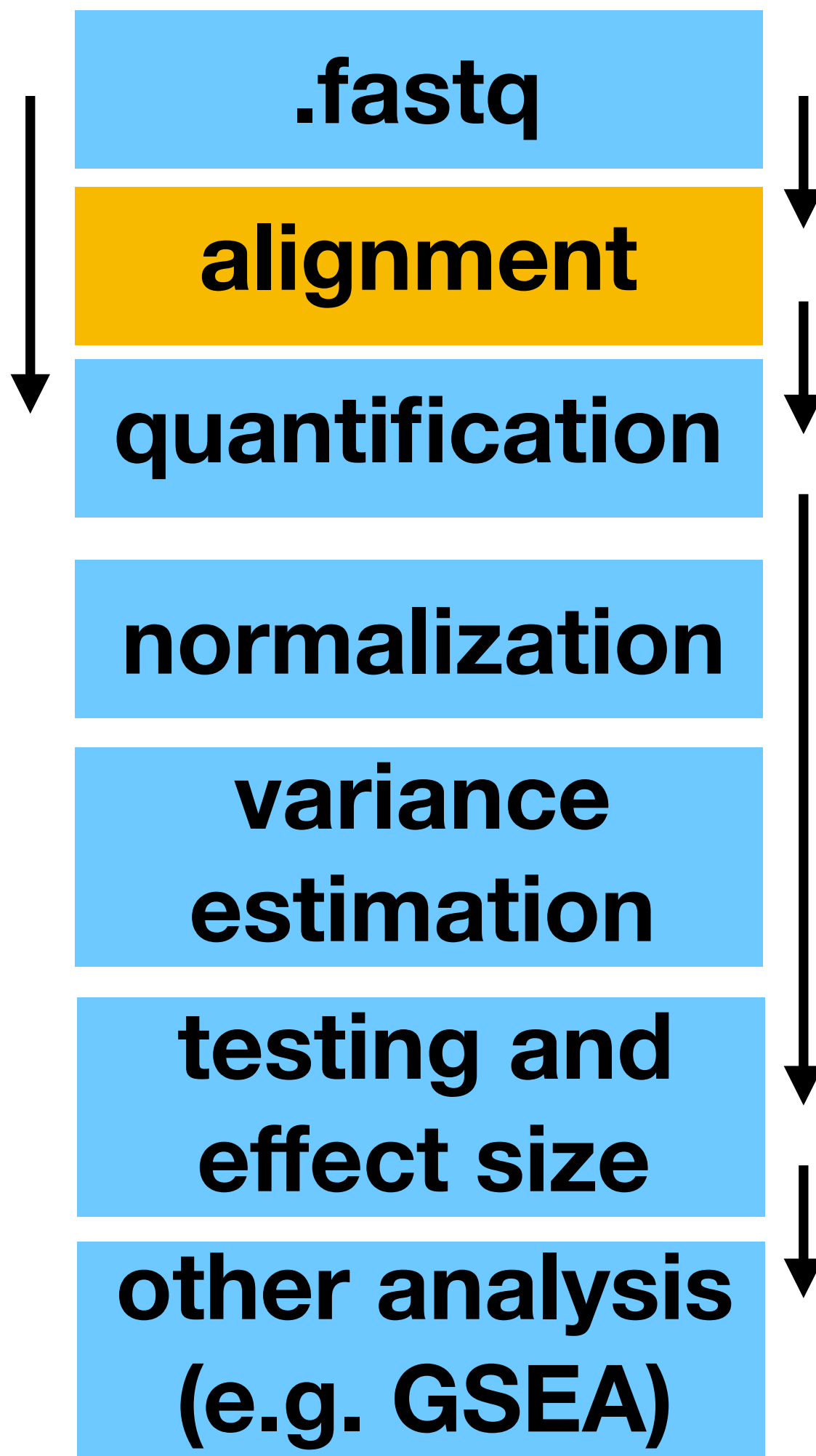
	In gene set	Not in gene set	
DGE	a	b	a + b
Background	c	d	c + d
	a + c	b + d	

$$p = \frac{\binom{a+b}{a} \binom{c+d}{c}}{\binom{n}{a+c}} = \frac{(a+b)! (c+d)! (a+c)! (b+d)!}{a! b! c! d! n!}$$

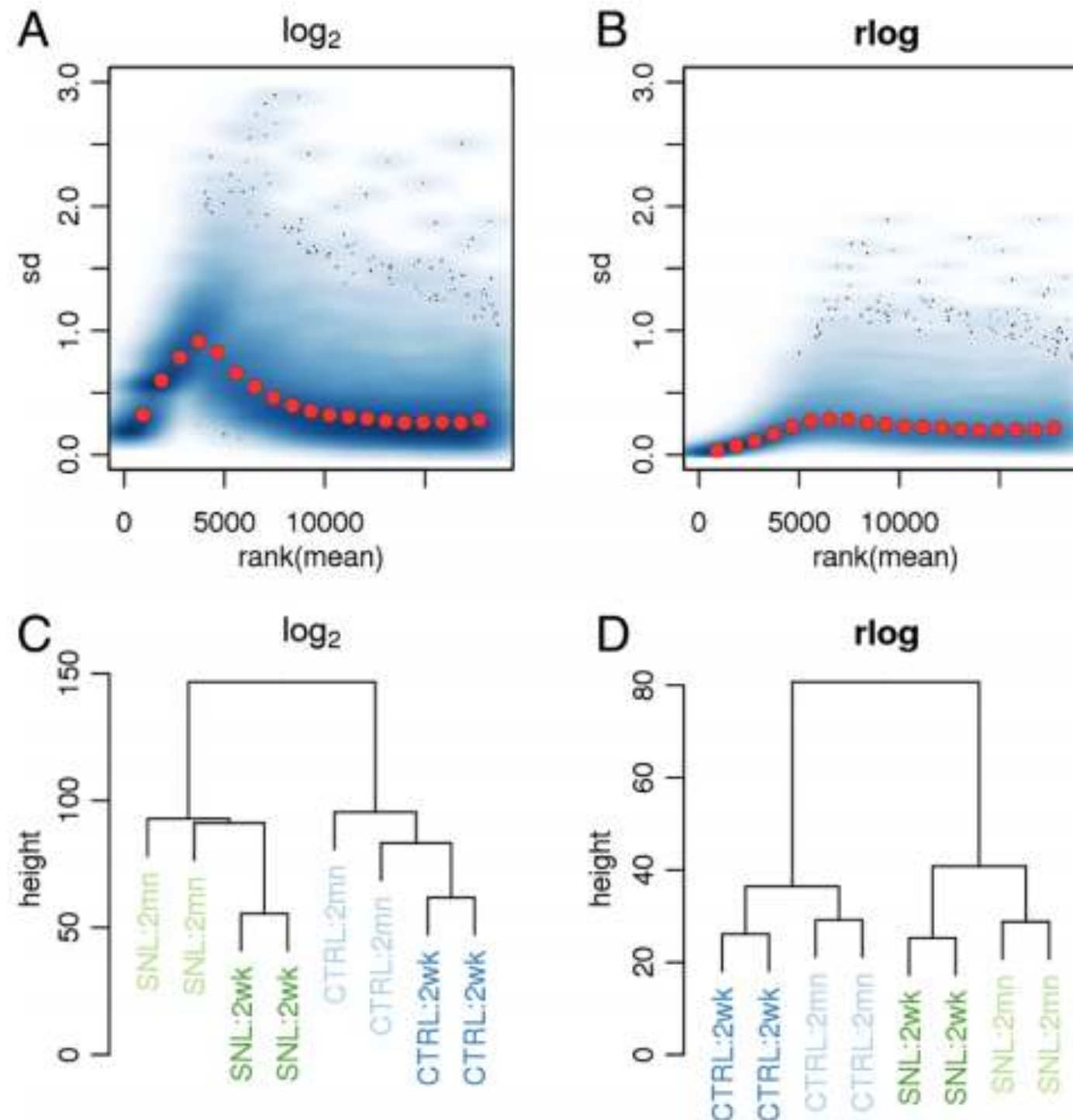
Fisher’s test

goseq estimates and corrects for RNA-seq biases in GSEA

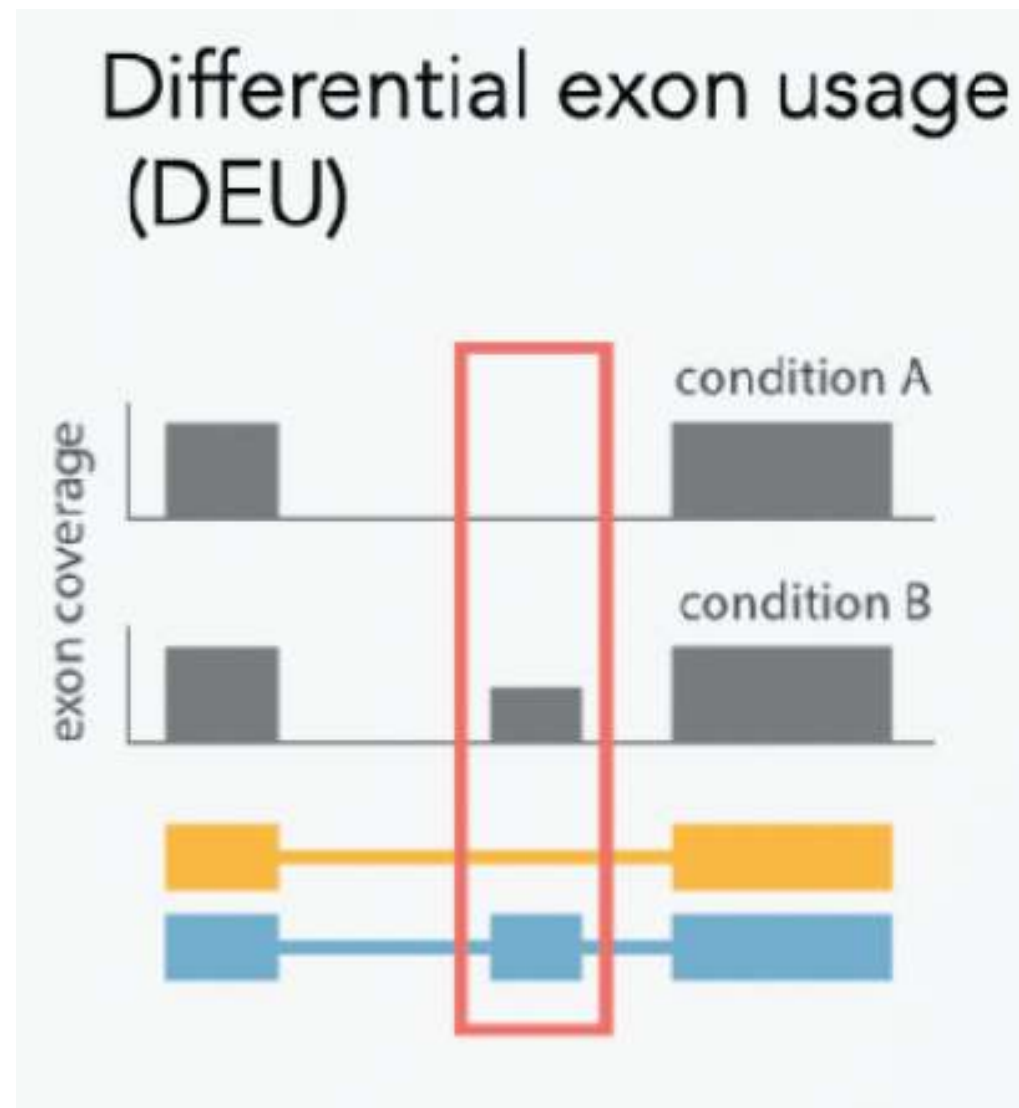




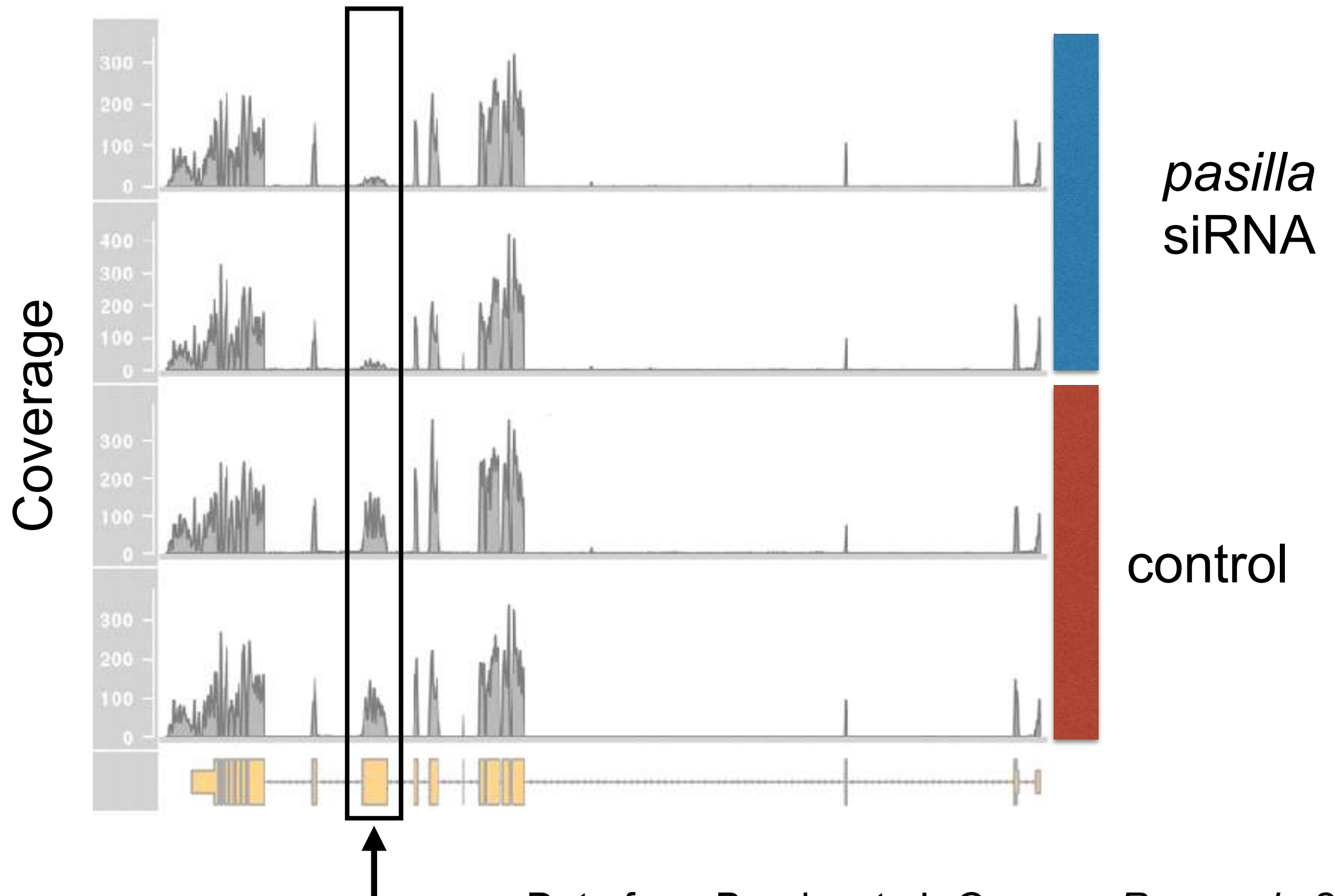
For exploratory analysis, clustering, PCA: Use rlog or vst data!



Differential exon usage



High-throughput RNA sequencing enables an unbiased characterisation of isoform expression



DEXSeq: inference of differential exon usage

samples →

exons ↓

	treated1fb	treated2fb	treated3fb	untreated1fb	untreated2fb	untreated3fb	untreated4fb
E001	1997	494	562	1150	2514	570	547
E002	122	112	180	69	203	156	142
E003	276	293	305	190	398	312	259
E004	420	200	182	230	446	183	185
E005	416	217	279	146	170	237	231
E006	486	357	471	190	337	418	364
E007	574	465	536	469	805	480	496
E008	536	417	447	541	832	475	472
E009	191	237	216	217	427	286	222
E010	188	130	96	617	1177	520	508
E011	165	212	210	118	275	294	269
E012	536	437	414	441	792	619	504
E013	72	41	49	40	76	34	38
E014	3	0	33	5	0	2	42

$$\text{exon usage} = \frac{\text{\# of transcripts including an exon}}{\text{\# total transcripts}}$$

DEXSeq: inference of differential exon usage

exon i , sample j

$l = 1$ exon under consideration

$l = 0$ sum of gene count

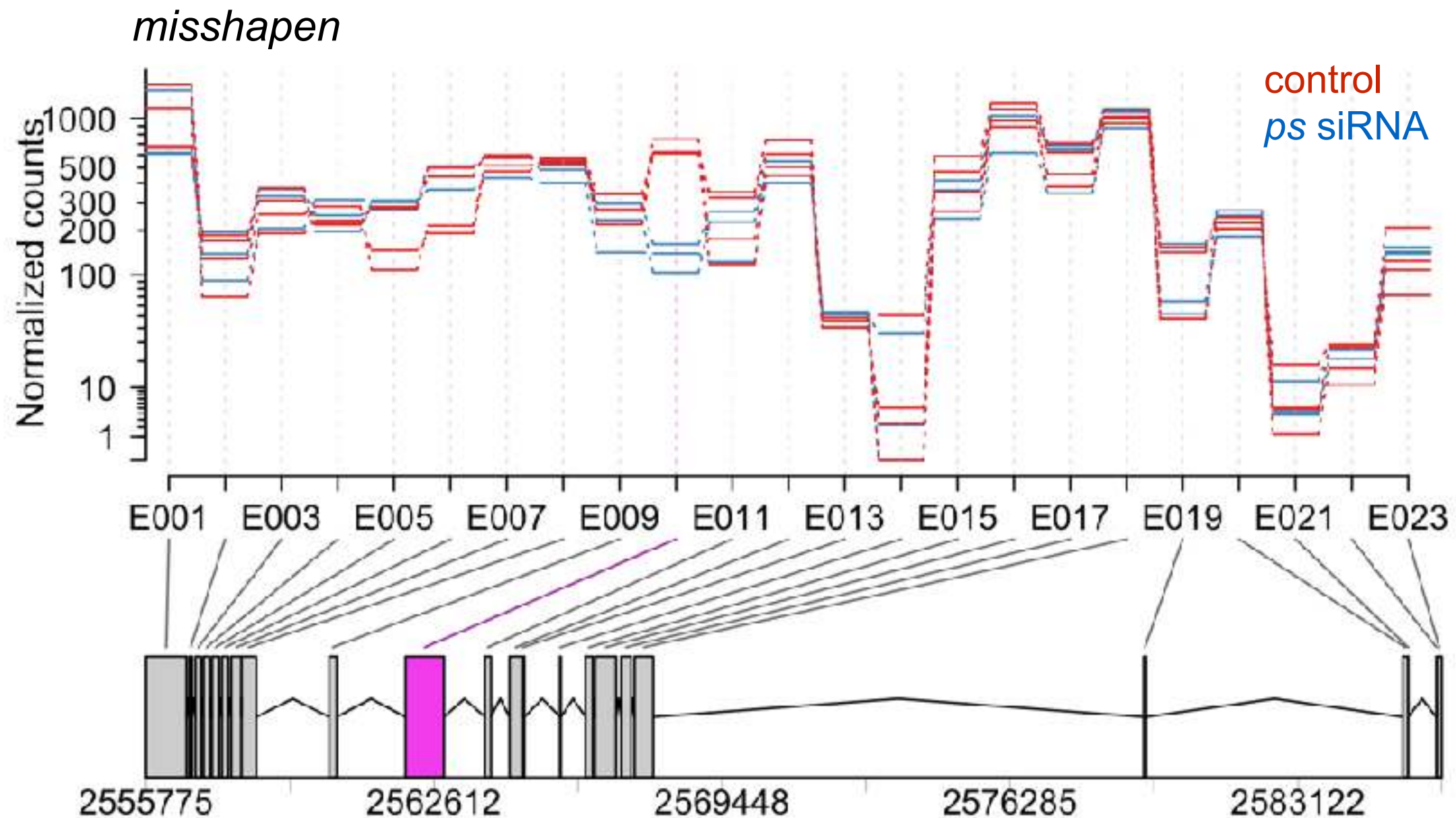
$$\log \mu_{ijl} = \beta_{ij}^S + l\beta_i^E + \beta_{i\rho_j}^{EC}$$

sample specific
contributions
(gene expression)

average
exon usage

changes
in exon usage
due to the
conditions

DEXSeq: inference of differential exon usage



Detecting differential splicing vs differential usage of TSS and polyA

